More on functional forms

EC 339

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Motivation

New functional forms

There is more to OLS than **linear-in-variables** models or **log-transformed** models.

But do these models **preserve** OLS *Classical Assumptions*?

- They do!
- But under what conditions?

As long as the model remains **linear in parameters**, everything is fine.

- 1. Regression through the **origin**
- 2. Regression with **quadratic** terms
- 3. Inverse forms
- 4. Interaction terms
- 5. Binary (dummy) variables

Regression through the origin

It is used whenever we need to impose the **restriction** that, when x = 0, the expected value of y is also zero.

It should be applied **only** when theory recommends to do so.

$$y_i = eta_1 x_{1i} + u_i$$

Regression through the origin

 $Cons_i = eta_1 Inc_i + u_i$



Many times, the effect of a variable x_i on y also depends on the **level** of that independent variable.

We can also apply quadratic terms when the effect of x_i on y changes after a given threshold.

$$y_i=eta_0+eta_1x_{1i}+eta_2(x_{1i})^2+\cdots+eta_kx_{ki}+u_i$$





$$wage_i = eta_0 + eta_1 exper_i + eta_2 exper_i^2 + u_i$$

Hourly wages vs. years of experience



$$wage_i = eta_0 + eta_1 exper_i + eta_2 exper_i^2 + u_i$$

Hourly wages vs. years of experience Hourly Wages (\$) 52 50 50



$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$



$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$



Interpretation

$$y_i=eta_0+eta_1x_{1i}+eta_2x_{1i}^2+u_i$$

$$rac{\partial \ y}{\partial \ x_1} = eta_1 + 2 \ \cdot \ eta_2 \ \cdot \ x_1$$

$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$

$$rac{\partial \ wage}{\partial \ educ} = eta_1 + 2 \ \cdot \ eta_2 \ \cdot \ educ$$

Inverse forms are used whenever the effect of an independent variable on y_i is expected to approach **zero** as its value approaches **infinity**.

As always, but especially important to this category, **economic theory** should *strongly recommend* the use of such functional form.

$$qchicken_i = eta_0 + eta_1 rac{1}{pchicken_i} + u_i$$



Chicken consumption vs. price of chicken

$$qchicken_i = eta_0 + eta_1 rac{1}{pchicken_i} + u_i$$



Chicken consumption vs. price of chicken

Interpretation

$$egin{aligned} y_i &= eta_0 + eta_1 rac{1}{x_{1i}} + u_i \ & \ rac{\partial \ y}{\partial \ x_1} = rac{-eta_1}{x_1^2} \end{aligned}$$

$$qchicken_i = eta_0 + eta_1 rac{1}{pchicken_i} + u_i
onumber \ rac{\partial \ qchicken}{\partial \ pchicken} = rac{-eta_1}{pchicken^2}$$

Interaction terms

Interaction terms

Whenever the effect of one variable on *y* depends on the **level of another variable**, the best **modeling strategy** is to use *interaction terms*.

For example, do we believe that an individual's **wage** depends on their **education**?

• If so, is this effect the **same** or **different** for two individuals with, e.g., a *college* degree, but with different years of experience on the job market?

• Then, we represent a model by

$$wage_i = eta_0 + eta_1 educ_i + eta_2 exper_i + eta_3 educ_i \cdot exper_i + u_i$$

In more general terms, regression estimates $(\hat{\beta}_i)$ describe **average effects**.

Some of these average effects may "hide" **heterogeneous effects** that differ by **group** or by the **level of another variable**.

Interaction terms help us in modeling such **heterogeneous** effects.

• For instance, it is plausible to consider that returns on education will differ by *gender*, *race*, *region*, etc.

Interaction terms

Interpretation

$$wage_i = eta_0 + eta_1 educ_i + eta_2 exper_i + eta_3 educ_i \cdot exper_i + u_i$$

$$rac{\partial \ wage}{\partial \ exper} = eta_2 + eta_3 \ \cdot \ educ$$

Binary variables

Categorical variables are used to translate **qualitative information** into **numbers**.

• For instance, race, gender, being employed or not, enrolled in EC 339 or not, etc.

The **easiest** way to work with qualitative information is by using **binary** (*dummy*) variables.

For example,

$$y_i=eta_0+eta_1D_i+u_i$$

where $D_i = 1$ if the criterion is fulfilled, and $D_i = 0$ otherwise.

Binary variables

When **interpreting** regression coefficients associated with *dummy* variables, the *intercept*'s interpretation changes slightly.

Moreover, the **slope** coefficient on D_i is not interpreted in the same way we are used to.

Consider:

 $interviews_i = eta_0 + eta_1 graduate_i + u_i$

where

- *interviews*_i is the number of interviews a candidate is called for in a given period;
- *graduate_i* equals 1 if she has graduated from college, and 0 otherwise.

Binary variables

 $interviews_i = eta_0 + eta_1 graduate_i + u_i$

For this model,

- β_0 is the expected number of interviews when $graduate_i = 0$ (non-graduates);
- β_1 is the expected **difference** in interview calls between graduates and non-graduates;
- And $\beta_0 + \beta_1$ is the expected number of interviews for graduates (when $graduate_i = 1$).

• In this case, non-graduates are the reference group.

 $interviews_i = eta_0 + eta_1 graduate_i + u_i$

The model above is an example of an **intercept** *dummy* variable.

• We only have different **intercepts** when comparing two groups, but **slopes** are the same.

In order to allow for different **slopes**, we appeal to interaction terms involving categorical variables

• i.e., **slope** *dummy* variables.

Log-Level Model

Important! If you have a **log-linear** model with a *binary* variable, the interpretation of the coefficient on that variable **changes**.

 $\log(y_i)=eta_0+eta_1D_i+u_i$

with *D* being a *dummy* variable.

Interpretation of β_1 :

- When D=1, y will increase by $100 imes \left(e^{eta_1}-1
 ight)$ percent.
- When D=0, y will decrease by $100 imes \left(e^{-eta_1}-1
 ight)$ percent.

Log-Level Example

Binary explanatory variable: inlf

- inlf = 1 if the i^{th} individual is in the labor force.
- inlf = 0 if the i^{th} individual is not in the labor force.

 $\widehat{log(sleep_i)} = 8.08 - 0.00365~inlf_i$

- How do we interpret the coefficient on inlf?
 - Labor force participants sleep 36.65% less than non-participants.
 - Individuals that are not in the labor force sleep 36.92% % more than participants.

Slope *dummy* variables

Hourly wages vs. years of education (by gender)

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Female=1, Non-female=0

Gender 🗕 0 🔺 1

Slope *dummy* variables

Hourly wages vs. years of education (by gender)



Female=1, Non-female=0

Gender 🔶 0 📥 1

Slope *dummy* variables

Interpretation

$$egin{aligned} y_i &= eta_0 + eta_1 x_{1i} + eta_2 D_i + eta_3 D_i x_{1i} + u_i \ & \ & rac{\partial \ y}{\partial \ x_1} = eta_1 + eta_3 \ \cdot \ D \ & \ & \ & rac{\partial \ y}{\partial \ D} = eta_2 + eta_3 \ \cdot \ x_1 \end{aligned}$$

 $wage_i = eta_0 + eta_1 educ_i + eta_2 female_i + eta_3 educ_i \cdot female_i + u_i$

$$rac{\partial \ wage}{\partial \ educ} = eta_1 + eta_3 \ \cdot \ female$$

$$rac{\partial \ wage}{\partial \ female} = eta_2 + eta_3 \ \cdot \ educ$$

Next time: Functional forms in practice