Regression Stuff EC 607, Set 05

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Prologue

Last time: Inference and simulation

Let's review using a quote from MHE

We've chosen to start with the **asymptotic approach to inference** because modern empirical work typically leans heavily on the large-sample theory that lies behind robust variance formulas. The **payoff is valid inference under weak assumptions**, in particular, a framework that makes sense for our less-than-literal approach to regression models. On the other hand, the **largesample approach is not without its dangers**...

MHE, p. 48 (emphasis added)

Schedule

Today

Regression and causality Read MHE 3.2

Upcoming

Project, step 1 Assignment #2 Advice Make sure you're taking a few minutes for personal health.⁺

t health = physical, mental, and spiritual. Also: Do a better job than I do.

Regression talk Saturated models

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 $Wages_i = \alpha + \beta \mathbb{I}\{College Graduate\}_i + \varepsilon_i$

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 $\operatorname{Wages}_i = lpha + eta_1 \, \mathbb{I}\{s_i = 1\}_i + eta_2 \, \mathbb{I}\{s_i = 2\}_i + \dots + eta_T \, \mathbb{I}\{s_i = T\}_i + arepsilon_i\}_i$

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Here, $s_i = 0$ is our reference level; β_j is the effect of j years of schooling.

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Here, $s_i=0$ is our reference level; eta_j is the effect of j years of schooling.

 $E[Wages_i \mid s_i = j] - E[Wages_i \mid s_i = 0] = \alpha + \beta_j - \alpha = \beta_j$

Saturated models

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A **Saturated models perfectly fit the CEF** because the CEF is a linear function of the dummy variables—a special case of the linear CEF theorem.

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Here, the uninteracted terms $(\beta_1 \& \beta_2)$ are called **main effects**; β_3 gives the effect of the **interaction**.

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 $egin{aligned} E[ext{Wages}_i | ext{College Graduate}_i &= 0, ext{ Female}_i &= 0] &= lpha \ E[ext{Wages}_i | ext{College Graduate}_i &= 1, ext{ Female}_i &= 0] &= lpha + eta_1 \ E[ext{Wages}_i | ext{College Graduate}_i &= 0, ext{ Female}_i &= 1] &= lpha + eta_2 \ E[ext{Wages}_i | ext{College Graduate}_i &= 1, ext{ Female}_i &= 1] &= lpha + eta_1 + eta_2 + eta_3 \end{aligned}$

Saturated models

The CEF can take on four possible values,

 $egin{aligned} &E[\mathrm{Wages}_i|\mathrm{College}\ \mathrm{Graduate}_i=0,\ \mathrm{Female}_i=0]=lpha\ &E[\mathrm{Wages}_i|\mathrm{College}\ \mathrm{Graduate}_i=1,\ \mathrm{Female}_i=0]=lpha+eta_1\ &E[\mathrm{Wages}_i|\mathrm{College}\ \mathrm{Graduate}_i=0,\ \mathrm{Female}_i=1]=lpha+eta_2\ &E[\mathrm{Wages}_i|\mathrm{College}\ \mathrm{Graduate}_i=1,\ \mathrm{Female}_i=1]=lpha+eta_1+eta_2+eta_3 \end{aligned}$

and the specification of our saturated regression model

 $egin{aligned} \operatorname{Wages}_i &= lpha + eta_1 \, \mathbb{I}\{\operatorname{College}\, \operatorname{Graduate}\}_i + eta_2 \, \mathbb{I}\{\operatorname{Female}\}_i \ &+ eta_3 \, \mathbb{I}\{\operatorname{College}\, \operatorname{Graduate}\}_i imes \mathbb{I}\{\operatorname{Female}\}_i + arepsilon_i \end{aligned}$

does not restrict the CEF at all.

Model specification

Saturated models sit at one extreme of the model-specification spectrum, with *linear, uninteracted models* occupying the opposite extreme.

Saturated models

- Fit CEF (+)
- Complex (-)
 - Many dummies
 - Many interactions

Plain, linear models

- Linear approximations (–)
- Simple (+)

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Don't forget the there are many options in between—though some make less sense than others (*e.g.*, interactions without main effects).

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Continuous, linear probability, logged, non-negative—it works for all.

Now back to causality...

The return of causality

We've spent the last few lectures developing properties/understanding of (1) the CEF and (2) least-squares regression.

Let's return to our main goal of the course...

Q When can we actually interpret a regression as **causal**?[†]

† Hint: There is no "reg y x, causal" command in Stata.

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A A regression is causal when the CEF it approximates is causal.

The return of causality

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A First, return to the potential-outcomes framework, describing hypothetical outcomes.

A CEF is causal when it describes **differences in average potential outcomes** for a fixed reference population.

MHE, p. 52 (emphasis added)

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Let's work through this "definition" of causal CEFs with an example.

Causal CEFs

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Now we would like to extend this framework to

- 1. variables that take on **more than two values**
- 2. situations that require us to **hold many covariates constant** in order to achieve a valid causal interpretation

Causal CEFs

The idea of *holding (many) covariates constant* brings us to one of the cornerstones of applied econometrics: the **conditional independence assumption (CIA)** (also called *selection on observables*).

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To see how CIA eliminates selection bias...

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Even randomized experiments need the CIA—*e.g.*, the STAR experiment's *within-school* randomization.

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Now, Y_{si} denotes individual *i*'s outcome with *s* years of schooling.

Let each individual have her own function between schooling and earnings.

$$\mathrm{Y}_{si} \equiv f_i(s)$$
 .

 $f_i(s)$ answers exactly the type of causal questions that we want to answer.

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$$\begin{split} E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}, \, s_{i} = s] - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}, \, s_{i} = s - 1] \\ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_{i}, \, s_{i} = s] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}, \, s_{i} = s - 1] \\ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_{i}] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}] \\ &= E[\mathbf{Y}_{si} - \mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}] \\ &= E[f_{i}(s) - f_{i}(s - 1) \mid \mathbf{X}_{i}] \end{split}$$

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$$\begin{split} E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}, \, \boldsymbol{s}_{i} = \boldsymbol{s}] - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}, \, \boldsymbol{s}_{i} = \boldsymbol{s} - 1] \\ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_{i}, \, \boldsymbol{s}_{i} = \boldsymbol{s}] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}, \, \boldsymbol{s}_{i} = \boldsymbol{s} - 1] \\ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_{i}] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}] \\ &= E[\mathbf{Y}_{si} - \mathbf{Y}_{(s-1)i} \mid \mathbf{X}_{i}] \\ &= E[f_{i}(\boldsymbol{s}) - f_{i}(\boldsymbol{s} - 1) \mid \mathbf{X}_{i}] \end{split}$$

With the CIA, a difference in conditional averages allows causal interpretations.

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 $E[\mathbf{Y}_i \mid \mathbf{X}_i, \, \boldsymbol{s}_i = 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, \boldsymbol{s}_i = 11]$

The conditional independence assumption

Example The causal effect of high-school graduation is

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= The (conditional) average causal effect of graduation at X_i

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Now take the expected value of both sides and apply the LIE.

$$egin{aligned} & Eiggl(E[\mathbf{Y}_i \mid \mathbf{X}_i, \, oldsymbol{s_i} = \mathbf{12}] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, oldsymbol{s_i} = \mathbf{11}] iggr) \ & = Eiggl(E[f_i(\mathbf{12}) - f_i(\mathbf{11}) \mid \mathbf{X}_i] iggr) \end{aligned}$$

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Q What about the **unconditional** average causal effect of graduation?

A First, remember what we just showed...

$$E[\mathrm{Y}_i \mid \mathrm{X}_i, \, oldsymbol{s}_i = 12] - E[\mathrm{Y}_i \mid \mathrm{X}_i, \, oldsymbol{s}_i = 11] = E[f_i(12) - f_i(11) \mid \mathrm{X}_i]$$

Now take the expected value of both sides and apply the LIE.

$$E\left(E[\mathbf{Y}_i \mid \mathbf{X}_i, \, \mathbf{s}_i = \mathbf{12}] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, \mathbf{s}_i = \mathbf{11}]\right)$$
$$= E\left(E[f_i(\mathbf{12}) - f_i(\mathbf{11}) \mid \mathbf{X}_i]\right)$$
$$= E[f_i(\mathbf{12}) - f_i(\mathbf{11})] \quad (\text{Iterating expectations})$$

The conditional independence assumption

Takeaways

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Takeaways

- 1. Conditional independence gives our parameters **causal interpretations** (eliminating selection bias).
- 2. The interpretation changes slightly—without iterating expectations, we have **conditional average treatment effects**.
- 3. The CIA is challenging—you need to know which set of covariates (X_i) leads to **as-good-as-random residual variation in your treatment**.
- 4. The idea of conditioning on observables to match *comparable* individuals introduces us to **matching estimators**—comparing groups of individuals with the same covariate values.

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Let's start with the 'easier' case: a linear, constant-effects (causal) model.

From the CIA to regression

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From the CIA to regression

Let $f_i(s)$ be linear in s and equal across i except for an error term, e.g.,

$$f_i(s) = lpha +
ho s + \eta_i$$
 (A)

Substitute in our observed value of s_i and the outcome \mathbf{Y}_i

$$\mathbf{Y}_i = \alpha + \rho \boldsymbol{s_i} + \eta_i \tag{B}$$

While ρ in (A) is explicitly causal, regression-based estimates of ρ in (B) need not be causal (selection/OVB for endogenous s_i).

From the CIA to regression

Continuing with our linear, constant-effect causal model...

$$f_i(s) = \alpha + \rho s + \eta_i \tag{A}$$

Now impose the conditional independence assumption for covariates X_i .

$$\eta_i = \mathbf{X}'_i \gamma + \nu_i \tag{C}$$

where γ is a vector of population coefficients from regressing η_i on \mathbf{X}_i .

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Note Least-squares regression implies

1.
$$E[\eta_i \mid \mathbf{X}_i] = \mathbf{X}'_i \boldsymbol{\gamma}$$

2. \mathbf{X}_i is uncorrelated with ν_i .

From the CIA to regression

Now write out the conditional expectation function of $f_i(s)$ on X_i and s_i .

 $egin{aligned} & E[f_i(s) \mid \mathbf{X}_i, \, s_i] \ & = E[f_i(s) \mid \mathbf{X}_i] \quad (\end{tabular}$

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$$\begin{split} E[f_i(s) \mid \mathbf{X}_i, \, s_i] \\ &= E[f_i(s) \mid \mathbf{X}_i] \quad \text{(CIA)} \\ &= E[\alpha + \rho s_i + \eta_i \mid \mathbf{X}_i] \\ &= \alpha + \rho s_i + E[\eta_i \mid \mathbf{X}_i] \\ &= \alpha + \rho s_i + \mathbf{X}'_i \gamma \quad \text{(Least-squares regression)} \end{split}$$

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The CEF of $f_i(s_i)$ is linear, which means that the (right⁺) population regression will be the CEF.

 \dagger Here, "right" means conditional on X_i .

From the CIA to regression

Thus, the linear causal (regression) model is

$$\mathrm{Y}_i = lpha +
ho s_i + \mathrm{X}_i' \gamma +
u_i$$

The residual u_i is uncorrelated with

- 1. s_i (from the CIA)
- 2. \mathbf{X}_i (from defining γ via the regression of η on \mathbf{X}_i)

The coefficient ρ gives the causal effect of s_i on Y_i .

From the CIA to regression

As Angrist and Pischke note, this **conditional-independence assumption** (*a.k.a.* the selection-on-observables assumption) is the cornerstone of modern empirical work in economics—and many other disciplines.

Nearly any empirical application that wants a causal interpretation involves a (sometimes implicit) argument that **conditional on some set of covariates, treatment is as-good-as random**.

From the CIA to regression

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Part of our job: Reasoning through the validity of this assumption.

CIA example

Let's continue with the returns to graduation (G_i) .

Let's imagine

- 1. Women are more likely to graduate.
- 2. Everyone receives the same return to graduation.
- 3. Women receive lower wages across the board.

CIA example

First, we need to generate some data.

```
# Set seed
set.seed(12345)
# Set sample size
n = 1e4
# Generate data
ex_df = tibble(
   female = rep(c(0, 1), each = n/2),
   grad = runif(n, min = female/3, max = 1) %>% round(0),
   wage = 100 - 25 * female + 5 * grad + rnorm(n, sd = 3)
)
```

CIA example

Now we can estimate our naïve regression

 $\mathrm{Wage}_i = \alpha + \beta \mathrm{Grad}_i + \varepsilon_i$

CIA example

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$$\mathrm{Wage}_i = \alpha + \beta \mathrm{Grad}_i + \varepsilon_i$$

lm(wage ~ grad, data = ex_df)

Coef. S.E. t stat

Intercept 91.65 0.20 447.70 Graduate -1.59 0.26 -6.18

CIA example

Now we can estimate our naïve regression

$$\mathrm{Wage}_i = lpha + \beta \mathrm{Grad}_i + arepsilon_i$$

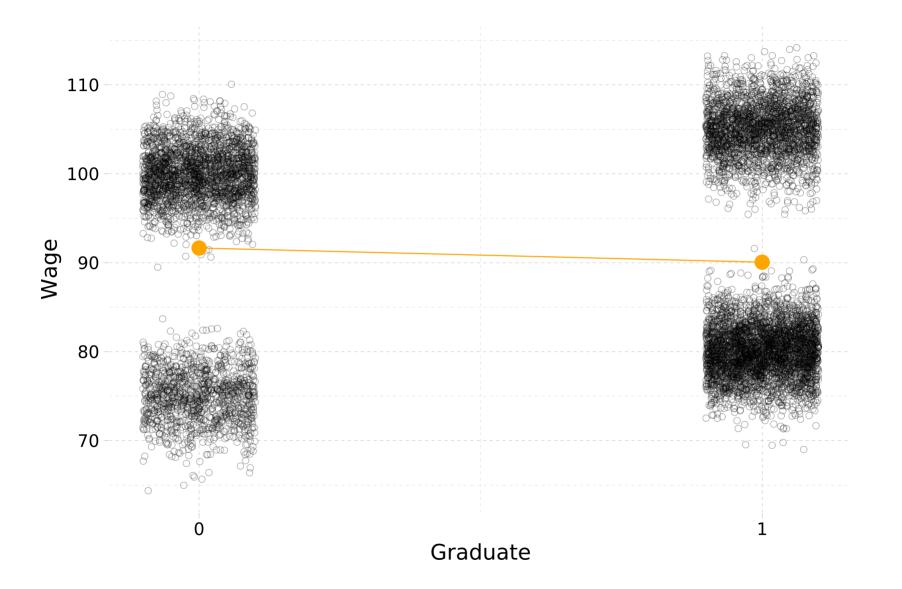
lm(wage ~ grad, data = ex_df)

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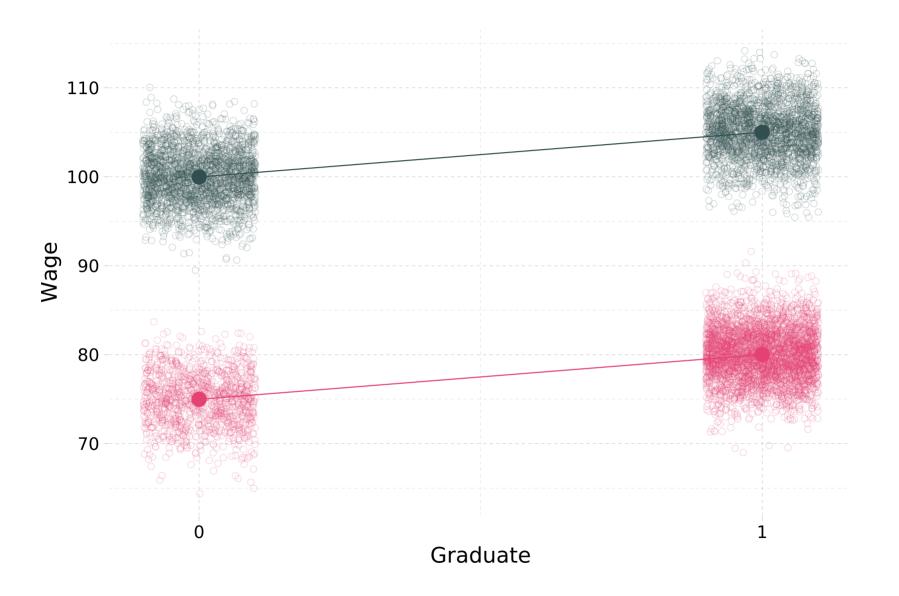
Intercept 91.65 0.20 447.70

Graduate -1.59 0.26 -6.18

Maybe we should have plotted our data...



We're still missing something...



CIA example

Now we can estimate our causal regression

 $\mathrm{Wage}_i = lpha + eta_1 \mathrm{Grad}_i + eta_2 \mathrm{Female}_i + arepsilon_i$

CIA example

Now we can estimate our causal regression

$$\operatorname{Wage}_i = \alpha + \beta_1 \operatorname{Grad}_i + \beta_2 \operatorname{Female}_i + \varepsilon_i$$

lm(wage ~ grad + female, data = ex_df)

	Coef.	S.E.	t stat
Intercept	99.98	0.05	1868.81
Graduate	5.03	0.06	78.23
Female	-25.00	0.06	-402.64

CIA example

Now we could also (unnecessarily) saturate the model...

 $\operatorname{Wage}_i = \alpha + \beta_1 \operatorname{Grad}_i + \beta_2 \operatorname{Female}_i + \beta_3 \operatorname{Grad}_i \times \operatorname{Female}_i + \varepsilon_i$

lm(wage ~ grad * female, data = ex_df)

	Coef.	S.E.	t stat
Intercept	99.98	0.06	1654.68
Graduate	5.03	0.08	59.28
Female	-24.99	0.10	-238.73
Graduate x Female	-0.01	0.13	-0.04

Summary

As always, assumptions matter.

- When is the CEF causal?
- Do you have a plausible/compelling argument for a valid CIA?

Least-squares regression helps estimate, but it also rests on assumptions.

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