Why Regression? EC 607, Set 03

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Prologue

Schedule

Last time

- The Experimental Ideal
- Fundamentals of $\ensuremath{\mathbb{R}}$

Today

What's so great about linear regression and OLS? **Read** *MHE* 3.1

Upcoming

Assignment₁ Custom OLS function fun.Assignment₂ First step of project proposal.

Why?

In our previous discussion, we began moving from simple differences to a regression framework.

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A As we discussed, regression allows us to control for covariates that *can* assist with (1) causal identification and (2) inference.

There's a deeper reason that we care about *linear* regression and ordinary least squares (OLS): *the conditional expectation function (CEF)*.

Why?

Even ignoring causality, we can show important relationships between

- 1. the CEF (the conditional expectation function),
- 2. the **population regression function**,
- 3. and the **sampling distribution of regression estimates**.

The CEF

Definition The **conditional expectation function** for a dependent variable Y_i , given a $K \times 1$ vector of covariates X_i , tells us the expected value (population average) of Y_i with X_i held constant.

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Examples

• $E[\text{Income}_i \mid \text{Education}_i]$

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- $E[\text{Income}_i \mid \text{Education}_i]$
- $E[Wage_i | Gender_i]$
- $E[\text{Birth weight}_i \mid \text{Air quality}_i]$

The CEF

Formally, for continuous Y_i with conditional density $f_y(t|\mathrm{X}_i=x)$,

$$E[\mathrm{Y}_i \mid \mathrm{X}_i = x] = \int t f_y(t | \mathrm{X}_i = x) dt$$

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and for discrete Y_i with conditional p.m.f. $\Pr(\mathrm{Y}_i = t | \mathrm{X}_i = x)$,

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Notice We are focusing on the **population**.

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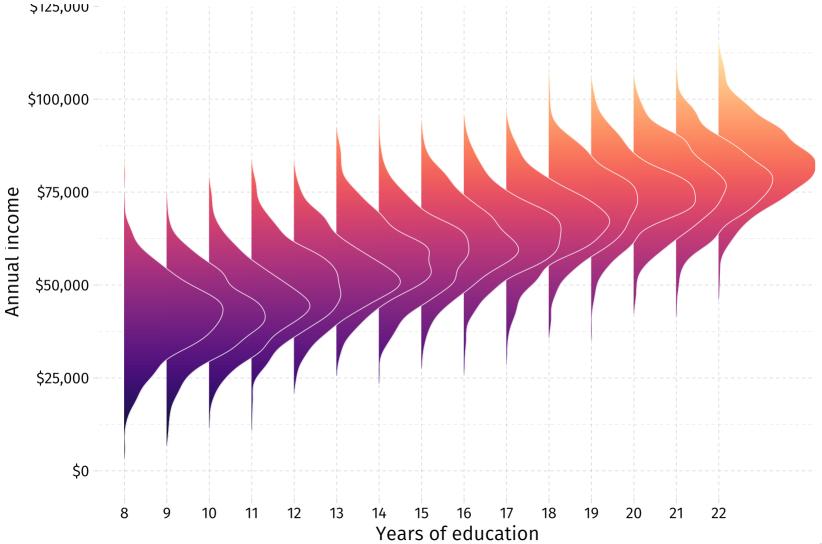
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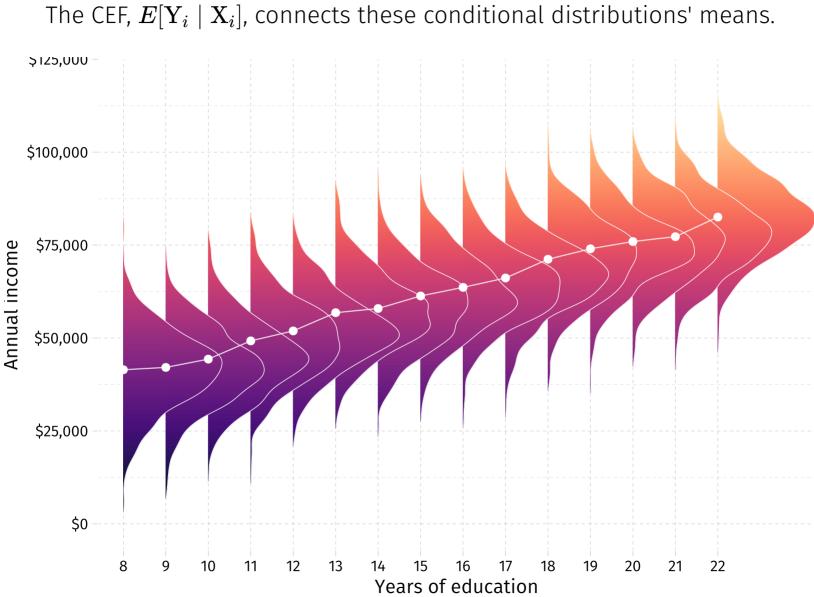
$$E[\mathrm{Y}_i \mid \mathrm{X}_i = x] = \sum_t t \Pr(\mathrm{Y}_i = t | \mathrm{X}_i = x)$$

Notice We are focusing on the **population**. We want to build our intuition about the parameters that we will eventually estimate.

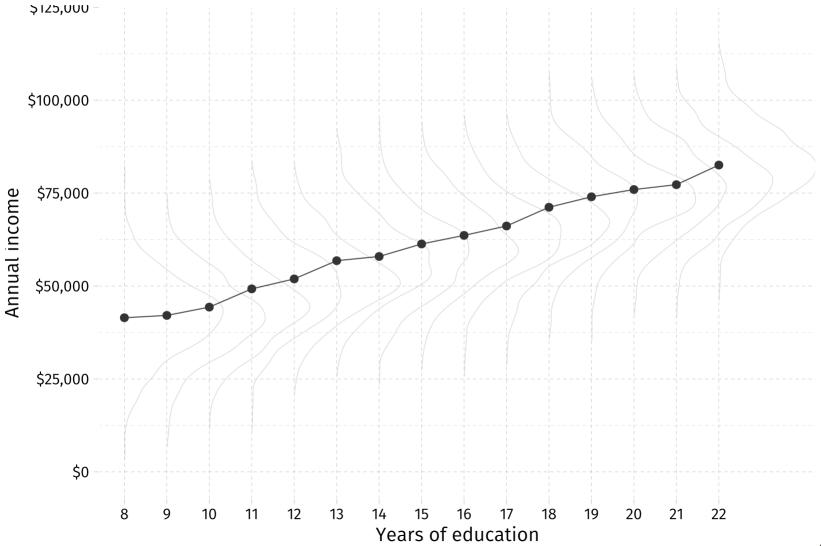
Graphically...

The conditional distributions of Y_i for $X_i = x$ in 8, ..., 22.





Focusing in on the CEF, $E[\mathbf{Y}_i \mid \mathbf{X}_i]$...



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Q How does the CEF relate to/inform regression?

The CEF

As we derive the properties and relationships associated with the CEF, regression, and a host of other estimators, we will frequently rely upon **the Law of Iterated Expectations** (LIE).

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The CEF

As we derive the properties and relationships associated with the CEF, regression, and a host of other estimators, we will frequently rely upon **the Law of Iterated Expectations** (LIE).

$$E[\mathbf{Y}_i] = E\left(\left. E[\mathbf{Y}_i \mid \mathbf{X}_i]
ight)$$

which says that the **unconditional expectation** is equal to the **unconditional average** of the **conditional expectation function**.

A proof of the LIE

First, we need notation...

Let $f_{x,y}(u,t)$ denote the joint density for continuous RVs (X_i, Y_i) .

Let $f_{y|x}(t \mid X_i = u)$ denote the conditional distribution of Y_i given $X_i = u$.

And let $g_y(t)$ and $g_x(u)$ denote the marginal densities of Y_i and X_i .

$$Eig(E[\mathbf{Y}_i \mid \mathbf{X}_i] ig)$$

$$egin{aligned} & Eig(egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i] ig) \ & = \int E[\mathbf{Y}_i \mid \mathbf{X}_i = u] \, g_x(u) du \end{aligned}$$

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$$egin{aligned} &Eiggl(egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i] iggr) \ &= \int E[\mathbf{Y}_i \mid \mathbf{X}_i = u] \, g_x(u) du \ &= \int iggl[\int t \, f_{y|x}(t \mid \mathbf{X}_i = u) \, dt iggr] \, g_x(u) du \, dt \ &= \int t \, f_{y|x}(t \mid \mathbf{X}_i = u) \, g_x(u) du \, dt \ &= \int t \, iggl[\int f_{y|x}(t \mid \mathbf{X}_i = u) \, g_x(u) du iggr] \, dt \ &= \int t \, iggl[\int f_{x,y}(u,t) du iggr] \, dt \ &= \int t \, g_y(t) \, dt \ &= E[\mathbf{Y}_i] \end{aligned}$$

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Great. What's the point?

The LIE and the CEF

Theorem The CEF decomposition property (3.1.1)

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The LIE allows us to **decompose random variables** into two pieces

 $\mathbf{Y}_i = E[\mathbf{Y}_i \mid \mathbf{X}_i] + \varepsilon_i$

- 1. the conditional expectation function
- 2. **a residual** with special powers[†]
 - i. ε_i is mean independent of \mathbf{X}_i , *i.e.*, $E[\varepsilon_i \mid \mathbf{X}_i] = 0$.
 - ii. ε_i is uncorrelated with any function of \mathbf{X}_i .

⁺ Angrist and Pischke go with *special properties*.

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Proof The CEF decomposition property (properties i. and ii. of ε_i)

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Mean independence, $E[\varepsilon_i \mid \mathbf{X}_i] = 0$ $E[\varepsilon_i \mid \mathbf{X}_i]$ $\mathbf{U} = E ig(\mathbf{Y}_i - oldsymbol{E} [\mathbf{Y}_i \mid \mathbf{X}_i] \Big| \mathbf{X}_i ig)$ $\mathbf{E} = E[\mathbf{Y}_i \mid \mathbf{X}_i] - Eig(\left. E[\mathbf{Y}_i \mid \mathbf{X}_i]
ight| \mathbf{X}_i ig) ig)$ $= E[\mathbf{Y}_i \mid \mathbf{X}_i] - E[\mathbf{Y}_i \mid \mathbf{X}_i]$ = 0

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The LIE and the CEF

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says that we can decompose any random variable (e.g., Y_i) into

- 1. a part that is explained by \mathbf{X}_i (*i.e.*, the CEF $E[\mathbf{Y}_i \mid \mathbf{X}_i]$),
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Why the CEF?

The CEF also presents an intuitive summary of the relationship between Y_i and X_i , since we are often use means to characterize random variables.

But (of course) there are more reasons to use the CEF...

The LIE and the CEF

Theorem The CEF prediction property (3.1.2)

Let $m(\mathbf{X}_i)$ be any function of \mathbf{X}_i . The CEF solves

$$egin{split} E[\mathbf{Y}_i \mid \mathbf{X}_i] = rgmin_{m(\mathbf{X}_i)} & Eigg[\left(\mathrm{Y}_i - m(\mathbf{X}_i)
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In other words, the **CEF** is the minimum mean-squared error (MMSE) predictor of Y_i given X_i .

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Notice

- 1. We haven't restricted m to any class of functions—it can be nonlinear.
- 2. We're talking about *prediction* (specifically predicting Y_i).

$$\left(\mathrm{Y}_{i} - m(\mathrm{X}_{i})
ight)^{2}$$

(1)

$$\begin{split} \left(\mathbf{Y}_i - m(\mathbf{X}_i)\right)^2 \\ &= \left(\left\{\mathbf{Y}_i - \boldsymbol{E}[\mathbf{Y}_i \mid \mathbf{X}_i]\right\} + \left\{\boldsymbol{E}[\mathbf{Y}_i \mid \mathbf{X}_i] - m(\mathbf{X}_i)\right\}\right)^2 \end{split}$$

(1)

$$\begin{pmatrix} \mathbf{Y}_{i} - m(\mathbf{X}_{i}) \end{pmatrix}^{2}$$

$$= \left(\{ \mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] \} + \{ E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - m(\mathbf{X}_{i}) \} \right)^{2}$$

$$= \left(\mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] \right)^{2}$$

$$+ 2\left(E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - m(\mathbf{X}_{i}) \right) \times \left(\mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] \right)$$

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$$(\mathbf{c})$$

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$$(\mathbf{c})$$

Recall: We want to choose the $m(X_i)$ that minimizes (1) in expectation.

$$\begin{pmatrix} \mathbf{Y}_{i} - m(\mathbf{X}_{i}) \end{pmatrix}^{2}$$

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Recall: We want to choose the $m(X_i)$ that minimizes (1) in expectation. (a) is irrelevant, *i.e.*, it does not depend upon $m(X_i)$.

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$$+ \left(E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - m(\mathbf{X}_{i}) \right)^{2}$$

$$(c)$$

Recall: We want to choose the m(X_i) that minimizes (1) in expectation.
(a) is irrelevant, *i.e.*, it does not depend upon m(X_i).
(b) equals zero in expectation: E[h(X_i) × ε_i] = 0.

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Recall: We want to choose the $m(X_i)$ that minimizes (1) in expectation.

- (a) is irrelevant, *i.e.*, it does not depend upon $m(X_i)$.
- (**b**) equals zero in expectation: $E[h(\mathbf{X}_i) \times \varepsilon_i] = 0$.
- (c) is minimized by $m(X_i) = E[Y_i \mid X_i]$, *i.e.*, when $m(X_i)$ is the CEF.

The LIE and the CEF

 \therefore the CEF is the function that minimizes the mean-squared error (MSE)

$$egin{split} E[\mathrm{Y}_i \mid \mathrm{X}_i] = rgmin_{m(\mathrm{X}_i)} \ Eigg[(\mathrm{Y}_i - m(\mathrm{X}_i))^2igg] \end{split}$$

The LIE and the CEF

One final property of the CEF (very similar to the decomposition property)

Theorem The ANOVA theorem (3.1.3)

 $\operatorname{Var}(\mathbf{Y}_i) = \operatorname{Var}(\boldsymbol{E}[\mathbf{Y}_i \mid \mathbf{X}_i]) + \boldsymbol{E}[\operatorname{Var}(\mathbf{Y}_i \mid \mathbf{X}_i)]$

which says that we can decompose the variance in \mathbf{Y}_i into

1. the variance in the CEF

2. the variance of the residual

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Example Decomposing wage variation into (**1**) variation explained by workers' characteristics and (**2**) unexplained (residual) variation

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Example Decomposing wage variation into (**1**) variation explained by workers' characteristics and (**2**) unexplained (residual) variation

The proof centers on the independence from the decomposition property of the CEF.

We now understand the CEF a bit better. But how does the CEF actually relate to regression?

The CEF and regression

We've discussed how the CEF summarizes empirical relationships.

Previously we discussed how regression provides simple empirical insights.

Let's link these two concepts.

The CEF and regression

Population least-squares regression

We will focus on β , the vector (a $K \times 1$ matrix) of population, least-squares regression coefficients, *i.e.*,

$$eta = rgmin_{b} \, E \Big[ig(\mathrm{Y}_{i} - \mathrm{X}_{i}^{\prime} b ig)^{2} \Big]$$

where b and \mathbf{X}_i are also $K \times 1$, and \mathbf{Y}_i is a scalar.

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where b and \mathbf{X}_i are also $K imes \mathbf{1}$, and \mathbf{Y}_i is a scalar.

Taking the first-order condition gives

$$Eig[\mathrm{X}_i\left(\mathrm{Y}_i-\mathrm{X}_i'b
ight)ig]=0$$

The CEF and regression

From the first-order condition

$$Eig[\mathrm{X}_i\left(\mathrm{Y}_i-\mathrm{X}_i'b
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we can solve for b. We've defined the optimum as β . Thus,

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The CEF and regression

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Note The first-order conditions tell us that our least-squares population regression residuals $(e_i = Y_i - X'_i\beta)$ are uncorrelated with X_i .

Anatomy

Our "new" result: $eta = E\left[\mathbf{X}_{i}\mathbf{X}_{i}'
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In **simple linear regression** (an intercept and one regressor x_i),

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For **multivariate regression**, the coefficient on the kth regressor x_{ki} is

$$eta_k = rac{ ext{Cov}(ext{Y}_i, \, ilde{x}_{ki})}{ ext{Var}(ilde{x}_{ki})}$$

where $ilde{x}_{ki}$ is the residual from a regression of x_{ki} on all other covariates.

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This alternative formulation of least-squares coefficients is quite powerful.

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Why?

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Why? This expression illustrates how each coefficient in a least-squares regression represents the bivariate slope coefficient after controlling for the other covariates.

Anatomy

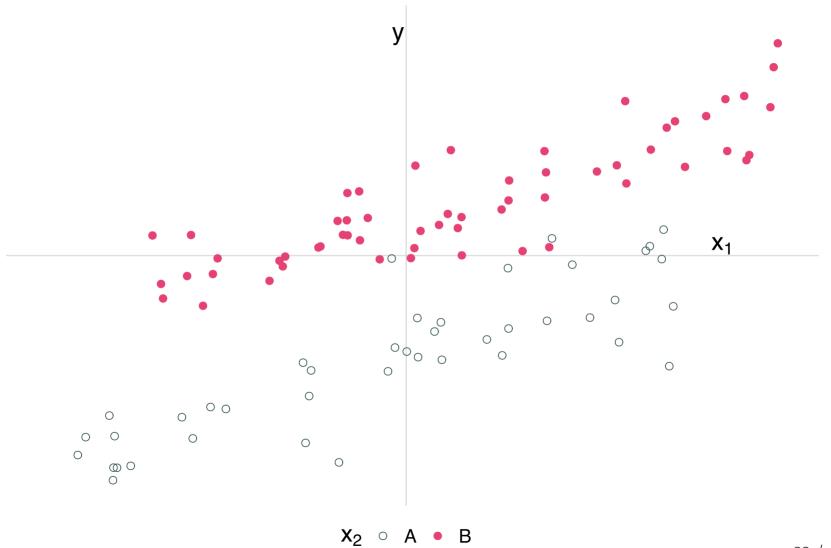
In fact, we can re-write our coefficients to further emphasize this point

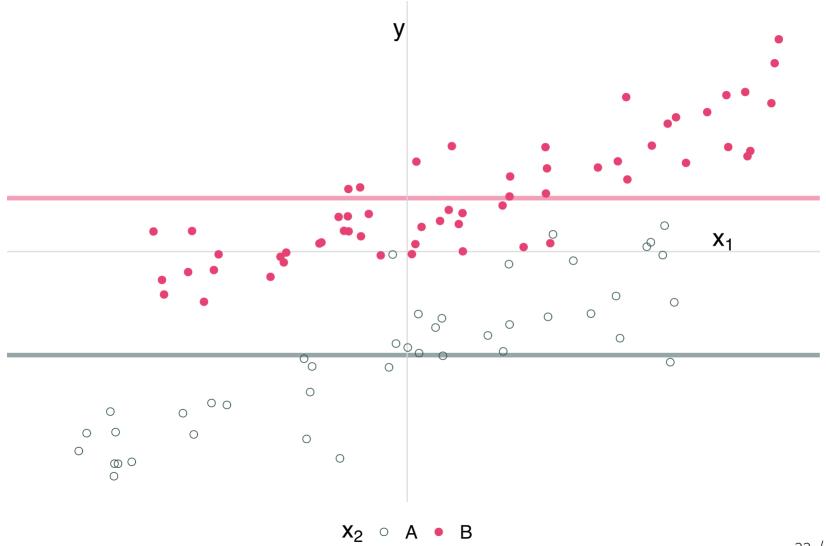
$$eta_k = rac{ ext{Cov} \Big(\widetilde{ ext{Y}}_i, \, ilde{x}_{ki} \Big) }{ ext{Var}(ilde{x}_{ki})}$$

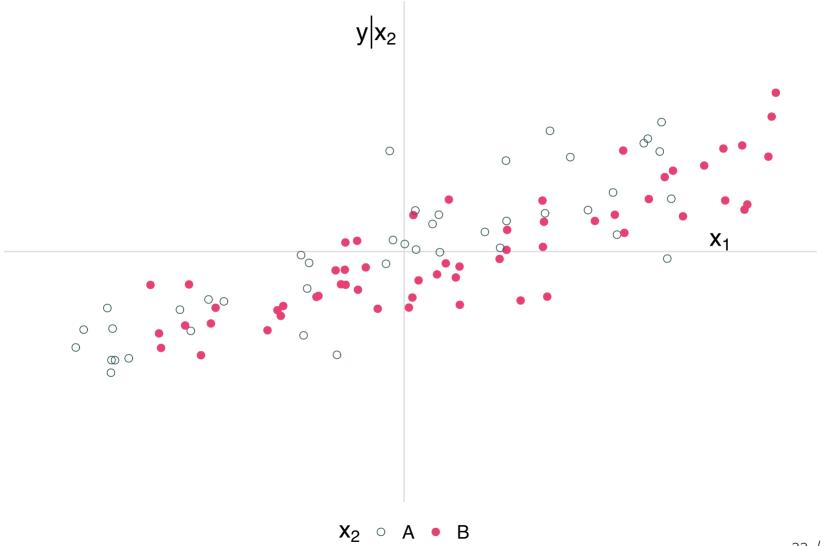
 $\widetilde{\mathrm{Y}}_i$ denotes the residual from regressing Y_i on all regressors except x_{ki} .

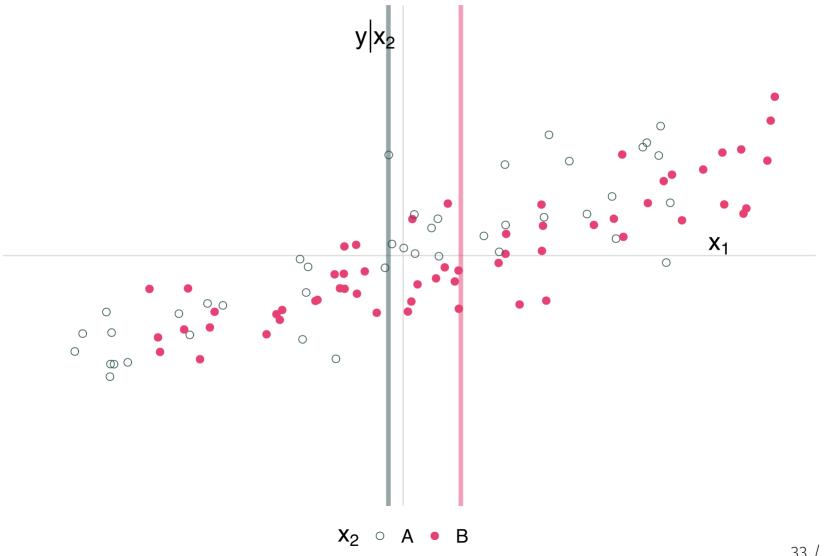
Graphical example

 $y_i=eta_0+eta_1x_{1i}+eta_2x_{2i}+arepsilon_i$

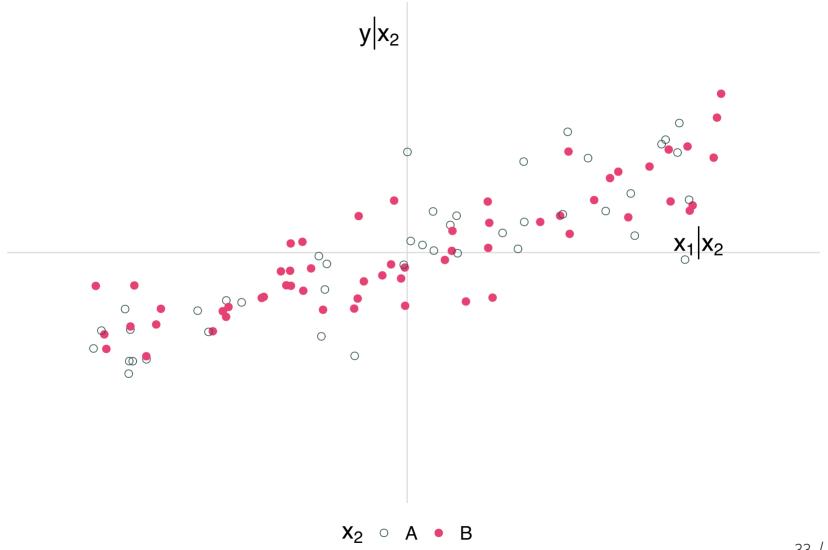


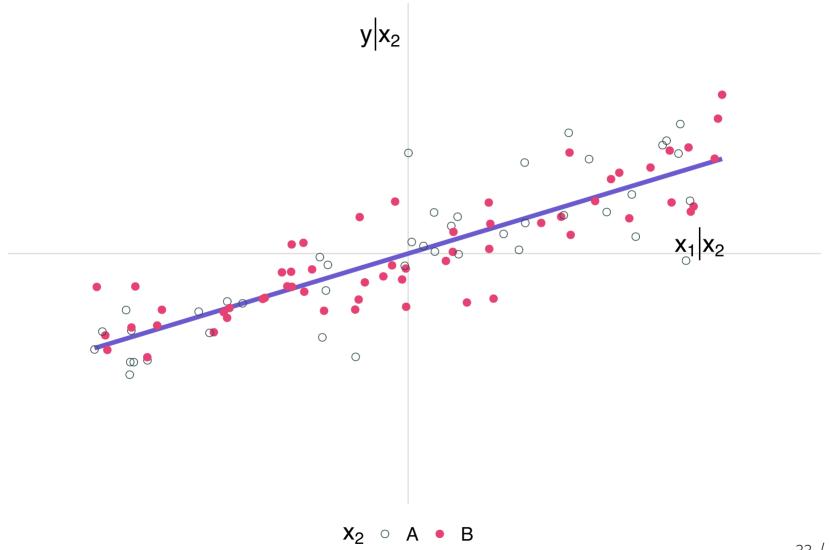






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Now that we've refreshed/deepened our regression knowledge, let's connect regression and the CEF.

Regression and the CEF

Angrist and Pischke make the case that

... you should be interested in regression parameters if you are interested in the CEF. (*MHE*, p.36)

Q What is the reasoning/connection?

Regression and the CEF

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Q What is the reasoning/connection?

A We'll cover three reasons.

1. *If the CEF is linear,* then the population regression line is the CEF.

2. The function $X'_i\beta$ is the min. MSE *linear* predictor of Y_i given X_i .

3. The function $\mathbf{X}'_i\beta$ gives the min. MSE *linear* approximation to the CEF.

Regression and the CEF

Theorem The linear CEF theorem (3.1.4)

If the CEF is linear, then the population regression is the CEF.

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Regression and the CEF

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Linearity can be a strong assumption. When might we expect linearity?

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Linearity can be a strong assumption. When might we expect linearity?

- 1. Situations in which (Y_i, X_i) follows a multivariate normal distribution. **Concern** Might be limited—especially when Y_i or X_i are not continuous.
- 2. Saturated regression models

Example A model with two binary indicators and their interaction.

Regression and the CEF

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 $X'_i\beta$ is the best *linear* predictor of Y_i given X_i (minimizes MSE).

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Proof We defined β as the vector that minimizes MSE, *i.e.*,

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so $\mathbf{X}'_{i}\beta$ is literally defined as the minimum MSE linear predictor of \mathbf{Y}_{i} .

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so $X'_i\beta$ is literally defined as the minimum MSE linear predictor of Y_i .

- The population-regression function $(X'_i\beta)$ is the best (min. MSE) *linear* predictor of Y_i given X_i .
- The CEF $(E[Y_i | X_i])$ is the best predictor (min. MSE) of Y_i given X_i across all classes of functions.

Regression and the CEF

Q If $X'_i\beta$ is **the best linear predictor** of Y_i given X_i , then why is there so much interest machine learning for prediction (opposed to regression)?

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A A few reasons

- 1. Relax linearity
- 2. Model selection
 - \circ choosing \mathbf{X}_i is not always obvious
 - overfitting is bad (bias-variance tradeoff)
- 3. It's fancy, shiny, and new
- 4. Some ML methods boil down to regression
- 5. Others?

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Counter Q Why are we (still) using regression?

Regression and the CEF

Theorem The regression CEF theorem (3.1.6)

The population regression function $X'_i\beta$ provides the minimum MSE linear approximation to the CEF $E[Y_i | X_i]$, *i.e.*,

$$eta = rgmin_b \, E iggl\{ iggl(E[\mathrm{Y}_i \mid \mathrm{X}_i] - \mathrm{X}_i'b iggr)^2 iggr\}$$

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Put simply Regression gives us the *best* linear approximation to the CEF.

$$\left(\mathbf{Y}_i - \mathbf{X}'_i b\right)^2$$
 (1)

$$\begin{aligned} \left(\mathbf{Y}_{i} - \mathbf{X}_{i}'b\right)^{2} \\ &= \left(\left\{\mathbf{Y}_{i} - \boldsymbol{E}[\mathbf{Y}_{i} \mid \mathbf{X}_{i}]\right\} + \left\{\boldsymbol{E}[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - \mathbf{X}_{i}'b\right\}\right)^{2} \end{aligned}$$
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We want to minimize (**b**), and we know β minimizes (**1**).

Proof First, recall that, in expectation, β is the *b* that minimizes $(Y_i - X'_i b)^2$

$$\begin{aligned} \left(\mathbf{Y}_{i} - \mathbf{X}_{i}^{\prime}b\right)^{2} & (\mathbf{1}) \\ &= \left(\left\{\mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}]\right\} + \left\{E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - \mathbf{X}_{i}^{\prime}b\right\}\right)^{2} & (\mathbf{a}) \\ &= \left(\mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}]\right)^{2} & (\mathbf{b}) \\ &+ \left(E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - \mathbf{X}_{i}^{\prime}b\right)^{2} & (\mathbf{b}) \\ &+ 2\left(\mathbf{Y}_{i} - E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}]\right)\left(E[\mathbf{Y}_{i} \mid \mathbf{X}_{i}] - \mathbf{X}_{i}^{\prime}b\right) & (\mathbf{c}) \end{aligned}$$

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- (a) is irrelevant, *i.e.*, it does not depend upon *b*.
- (c) can be written as $2\varepsilon_i h(\mathbf{X}_i)$, which equals zero in expectation.
- \therefore (In expectation) If $b = \beta$ minimizes (1), then $b = \beta$ minimizes (b).

Regression and the CEF

Let's review our new(-ish) regression results

- When the CEF is linear, the regression function is the CEF.
 Too small Very specific circumstances—or big assumptions.
- 2. Regression gives us the best *linear* predictor of Y_i (given X_i) **Off point** We're often interested in β -not \hat{Y}_i .
- Regression provides the best *linear* approximation of the CEF.
 Just right? (Depends on your goals)

Regression and the CEF

Motivation (3) tends to be the most compelling.

Even when the CEF is not linear, regression recovers the best linear approximation to the CEF.

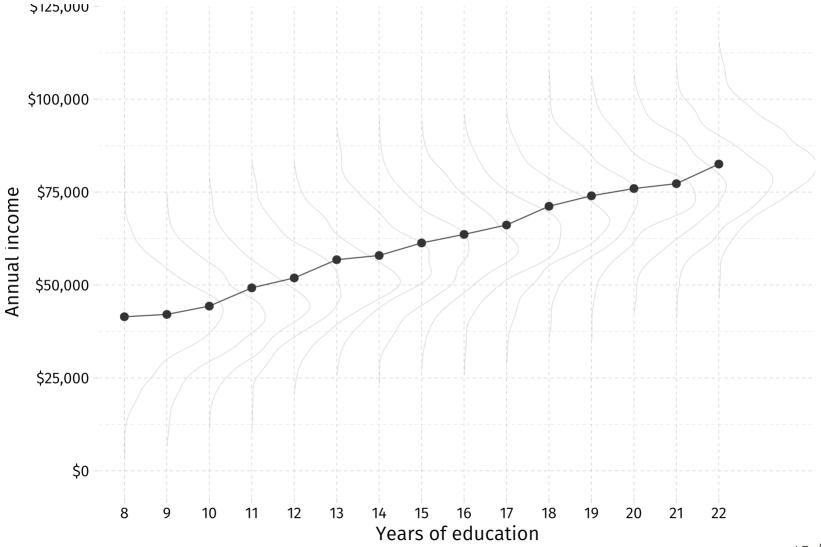
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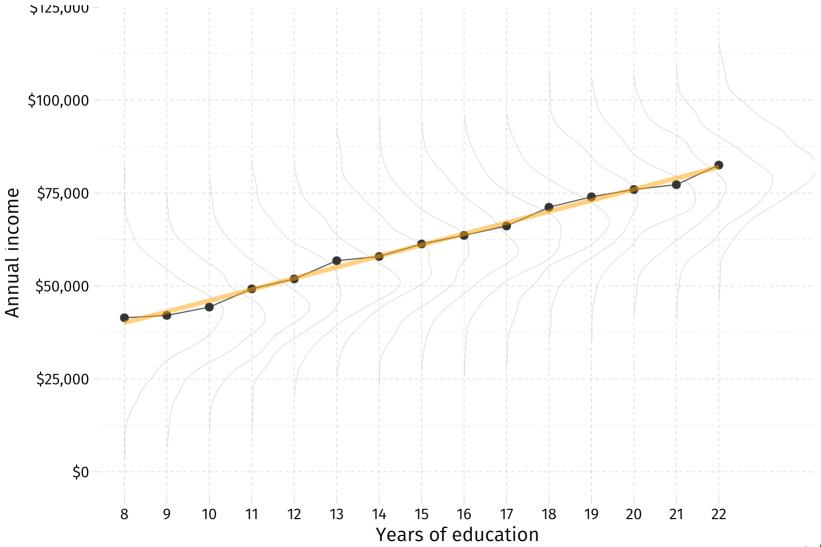
The statement that regression approximates the CEF lines up with our view of empirical work as an effort to describe the essential features of statistical relationships without necessarily trying to pin them down exactly. (*MHE*, p.39, emphasis added) Let's dig into this linear-approximate to the CEF a little more...

Returning to our **CEF**



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Adding the population regression function



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Regression and the CEF

As the previous figure suggests, one way to think about least-squares regression is **estimating a weighted regression on the CEF** rather than the individual observations.

Regression and the CEF

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TLDR Use $E[\mathbf{Y}_i | \mathbf{X}_i]$ as the outcome, rather than \mathbf{Y}_i , and properly weight.

Regression and the CEF

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TLDR Use $E[\mathbf{Y}_i \mid \mathbf{X}_i]$ as the outcome, rather than \mathbf{Y}_i , and properly weight.

Suppose X_i is discrete with pmf $g_x(u)$

$$Eigg[ig(egin{smallmatrix} E[\mathbf{Y}_i \mid \mathbf{X}_i] - \mathbf{X}_i'big)^2 igg] = \sum_uig(E[\mathbf{Y}_i \mid \mathbf{X}_i = u] - u'big)^2\,g_x(u)$$

i.e., β can be expressed as weighted-least squares regression of $E[Y_i \mid X_i = u]$ on u (the values of X_i) weighted by $g_x(u)$.

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Con You **will not** get the same standard errors.

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