# Inference and Simulation

EC 425/525, Set 4

Edward Rubin 17 April 2019

# Prologue

# Schedule

#### Last time

The CEF and least-squares regression

#### **Today**

Inference

Read MHE 3.1

#### **Upcoming**

Lab (as usual) on Friday. (Meet Jenni!)

No class on Monday.

Advice Don't ride elevators (especially in PLC).

# Why?

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A We rarely know the CEF or the population (and its regression vector).

We can draw statistical inferences about the population using samples.

Important The issue/topic of statistical inference is separate from causality.

Separate questions

- 1. How do we interpret the estimated coefficient  $\hat{\beta}$ ?
- 2. What is the sampling distribution of  $\hat{\beta}$ ?

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$$eta = E\left[ \mathrm{X}_i \mathrm{X}_i' 
ight]^{-1} E[\mathrm{X}_i \mathrm{Y}_i]$$

which we estimate via the ordinary least squares (OLS) estimator<sup>†</sup>

$$\hat{eta} = \left(\sum_i \mathrm{X}_i \mathrm{X}_i' 
ight)^{-1} \left(\sum_i \mathrm{X}_i \mathrm{Y}_i 
ight)^{-1}$$

† MHE presents a method-of-moments motivation for this derivation, where  $\frac{1}{n}\sum_i \mathbf{X}_i\mathbf{X}_i'$  is our sample-based estimated for  $E[\mathbf{X}_i\mathbf{X}_i']$ . You've also seen others, e.g., minimizing MSE of  $\mathbf{Y}_i$  given  $\mathbf{X}_i$ .

#### A classic

However you write it, this OLS estimator

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Note I'm following MHE in defining  $e_i = \mathrm{Y}_i - \mathrm{X}_i' \beta$ .

#### A classic

As you've learned, the OLS estimator

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has asymptotic covariance

$$E\left[\mathrm{X}_{i}\mathrm{X}_{i}^{\prime}
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which we estimate by (1) replacing  $e_i$  with  $\hat{e}_i = Y_i - X_i'\hat{\beta}$  and (2) replacing expectations with sample means, e.g.,  $E\left[X_iX_i'e_i^2\right]$  becomes  $\frac{1}{n}\sum\left[X_iX_i'\hat{e}_i^2\right]$ .

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Standard errors of this flavor are known as heteroskedasticity-consistent (or -robust) standard errors (or Eicker-Huber-White).

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Now, returning to to the asym. covariance matrix of  $\hat{\beta}$ ,

$$egin{aligned} E\left[\mathrm{X}_{i}\mathrm{X}_{i}^{\prime}
ight]^{-1}E\left[\mathrm{X}_{i}\mathrm{X}_{i}^{\prime}e_{i}^{2}
ight]E\left[\mathrm{X}_{i}\mathrm{X}_{i}^{\prime}
ight]^{-1} &= E\left[\mathrm{X}_{i}\mathrm{X}_{i}^{\prime}
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#### **Defaults**

Angrist and Pischke argue we should probably change our default to heteroskedasticity.

If the CEF is nonlinear, then our linear approximation (linear regression) generates heteroskedasticity.

$$egin{aligned} E\Big[ig(\mathbf{Y}_i - \mathbf{X}_i'etaig)^2 \mid \mathbf{X}_i\Big] \ &= E\Big[\Big(ig\{\mathbf{Y}_i - E[\mathbf{Y}_i \mid \mathbf{X}_i]ig\} + ig\{E[\mathbf{Y}_i \mid \mathbf{X}_i] - \mathbf{X}_i'etaig\}\Big)^2\Big|\mathbf{X}_i\Big] \ &= \mathrm{Var}(\mathbf{Y}_i \mid \mathbf{X}_i) + ig(E[\mathbf{Y}_i \mid \mathbf{X}_i] - \mathbf{X}_i'etaig)^2 \end{aligned}$$

Thus, even if  $Y_i \mid X_i$  has contant variance,  $e_i \mid X_i$  is heteroskedastic.

#### Two notes

1. Heteroskedasticity is **not our biggest concern** in inference.

...as an empirical matter, heteroskedasticity may matter very little... If heteroskedasticity matters a lot, say, more than a 30 percent increase or any marked decrease in standard errors, you should worry about possible programming errors or other problems. (MHE, p.47)

2. Notice that we've **avoided "standard" stronger assumptions**, e.g., normality, fixed regressors, linear CEF, homoskedasticity.

#### Two notes

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2. Notice that we've **avoided "standard" stronger assumptions**, e.g., normality, fixed regressors, linear CEF, homoskedasticity.

Following (2): We only have large-sample, asymptotic results (consistency) rather than finite-sample results (unbiasedness).

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One practical way we can study the behavior of an estimator: simulation.

Note You need to make sure your simulation can actually test/respond to the question you are asking (e.g., bias vs. consistency).

#### Simulation

Let's compare false- and true-positive rates<sup>†</sup> for

- 1. Homoskedasticity-assuming standard errors  $\left( \operatorname{Var}[e_i | \mathbf{X}_i] = \sigma^2 \right)$
- 2. Heteroskedasticity-robust standard errors

#### Simulation

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#### Simulation outline

- 1. Define data-generating process (DGP).
- 2. Choose sample size n.
- 3. Set seed.
- 4. Run 10,000 iterations of
  - a. Draw sample of size n from DGP.
  - b. Conduction inference.
  - c. Record inferences' outcomes.

<sup>†</sup> The false-positive rate goes by many names; another common name: type-I error rate.

# Data-generating process

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Let's keep the disturbances well behaved.

$$\mathbf{Y}_i = 1 + e^{0.5\mathbf{X}_i} + \varepsilon_i$$

where  $\mathrm{X}_i \sim \mathrm{Uniform}(0,10)$  and  $arepsilon_i \sim N(0,1)$ .

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```
library(pacman)

p_load(dplyr)

# Choose a size

n ← 1000

# Generate data

dgp_df ← tibble(

ε = rnorm(n, sd = 15),

x = runif(n, min = 0, max = 10),

y = 1 + exp(0.5 * x) + ε

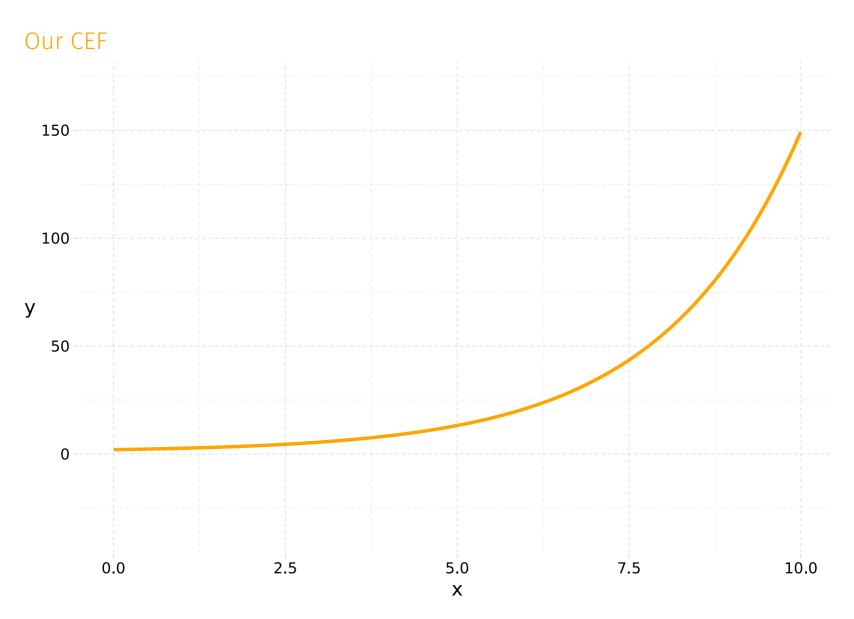
)
```

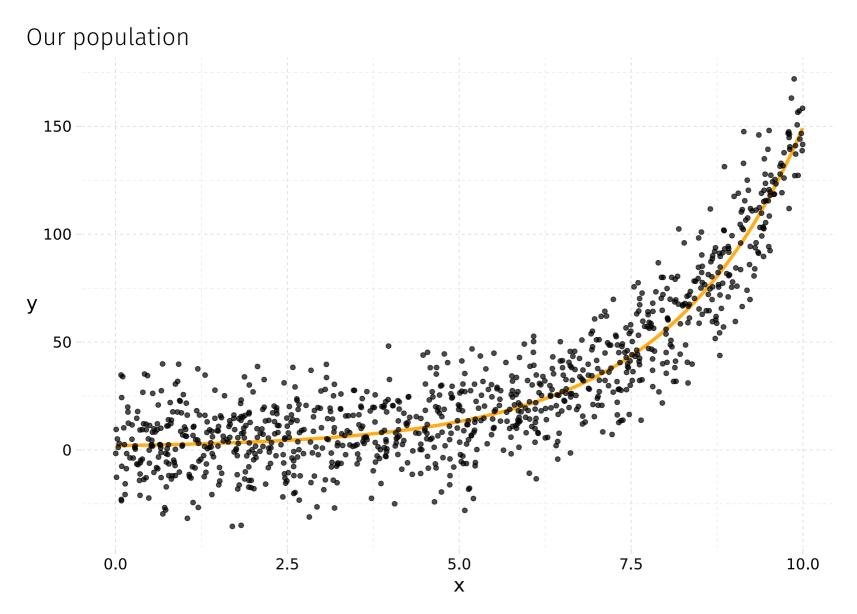
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```





# The population least-squares regression line 150 100 50

5.0 X

0.0

2.5

10.0

7.5

## Iterating

To make iterating easier, let's wrap our DGP in a function.

```
fun_iter ← function(iter, n = 30) {
    # Generate data
    iter_df ← tibble(
        ε = rnorm(n, sd = 15),
        x = runif(n, min = 0, max = 10),
        y = 1 + exp(0.5 * x) + ε
    )
}
```

We still need to run a regression and draw some inferences.

Note We're defaulting to size-30 samples.

We will use Im\_robust() from the estimatr package for OLS and inference.

- se\_type = "classical" provides homoskedasticity-assuming SEs
- se\_type = "HC2" provides heteroskedasticity-robust SEs

t lm() works for "spherical" standard errors but cannot calculate het.-robust standard errors.

### Inference

Now add these estimators to our iteration function...

```
fun iter \leftarrow function(iter, n = 30) {
  # Generate data
  iter df \leftarrow tibble(
    \varepsilon = rnorm(n, sd = 15),
    x = runif(n, min = 0, max = 10),
    y = 1 + \exp(0.5 * x) + \epsilon
  # Fstimate models
  lm1 \leftarrow lm \ robust(y \sim x, \ data = iter \ df, \ se \ type = "classical")
  lm2 \leftarrow lm \ robust(y \sim x, \ data = iter \ df, \ se \ type = "HC2")
  # Stack and return results
  bind rows(tidy(lm1), tidy(lm2)) %>%
    select(1:5) \%>\% filter(term = "x") \%>\%
    mutate(se type = c("classical", "HC2"), i = iter)
```

## Run it

Now we need to actually run our iter\_df() function 10,000 times.

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There are a lot of ways to run a single function over a list/vector of values.

- lapply(), e.g., lapply(X = 1:3, FUN = sqrt)
- for(), *e.g.*, for (x in 1:3) sqrt(x)
- map() from purrr, e.g., map(1:3, sqrt)

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- for(), *e.g.*, for (x in 1:3) sqrt(x)
- map() from purrr, e.g., map(1:3, sqrt)

We're going to go with map() from the purrr package because it easily parallelizes across platforms using the furrr package.

## Run it!

Run our function 10,000 times

```
# Packages
p_load(purrr)
# Set seed
set.seed(12345)
# Run 10,000 iterations
sim_list ← map(1:1e4, fun_iter)
```

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#### Parallelized 10,000 iterations

```
# Packages
p_load(purrr, furrr)
# Set options
set.seed(123)
# Tell R to parallelize
plan(multiprocess)
# Run 10,000 iterations
sim_list ← future_map(
    1:1e4, fun_iter,
    .options = future_options(seed = T)
)
```

## Run it!

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The furrr package (future + purrr) makes parallelization easy and fun!

## Run it!!

Our fun\_iter() function returns a data.frame, and future\_map() returns a list (of the returned objects).

So sim\_list is going to be a list of data.frame objects. We can bind them into one data.frame with bind\_rows().

```
# Bind list together
sim_df ← bind_rows(sim_list)
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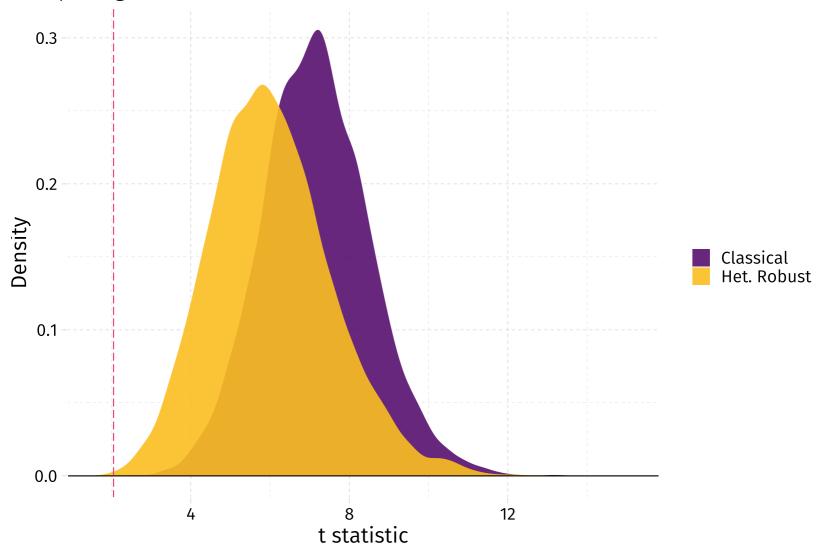
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sim_df ← bind_rows(sim_list)
```

So what are the results?

Comparing the distributions of standard errors for the coefficient on  $\boldsymbol{x}$ 



Comparing the distributions of t statistics for the coefficient on x



 $\mathbf{Q}$  All of these test are for a false  $\mathbf{H}_0$ . How would the simulation change to enforce a *true* null hypothesis?

## Updating to enforce the null

Let's update our simulation function to take arguments  $\gamma$  and  $\delta$  such that

$$\mathrm{Y}_i = 1 + e^{\gamma \mathrm{X}_i} + arepsilon_i$$

where  $arepsilon_i \sim \mathrm{N}(0, \sigma^2 \mathrm{X}_i^\delta)$ .

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where  $arepsilon_i \sim \mathrm{N}(0, \sigma^2 \mathrm{X}_i^\delta)$ .

In other words,

- $\gamma=0$  implies no relationship between  $\mathbf{Y}_i$  and  $\mathbf{X}_i$ .
- $\delta = 0$  implies homoskedasticity.

## Updating to enforce the null

Updating the function...

```
flex iter \leftarrow function(iter, y = 0, \delta = 1, n = 30) {
  # Generate data
  iter df \leftarrow tibble(
    x = runif(n, min = 0, max = 10),
     \varepsilon = \text{rnorm}(n, \text{sd} = 15 * x^{\delta}),
    v = 1 + \exp(v * x) + \varepsilon
  # Fstimate models
  lm1 \leftarrow lm \ robust(y \sim x, \ data = iter \ df, \ se \ type = "classical")
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## Run again!

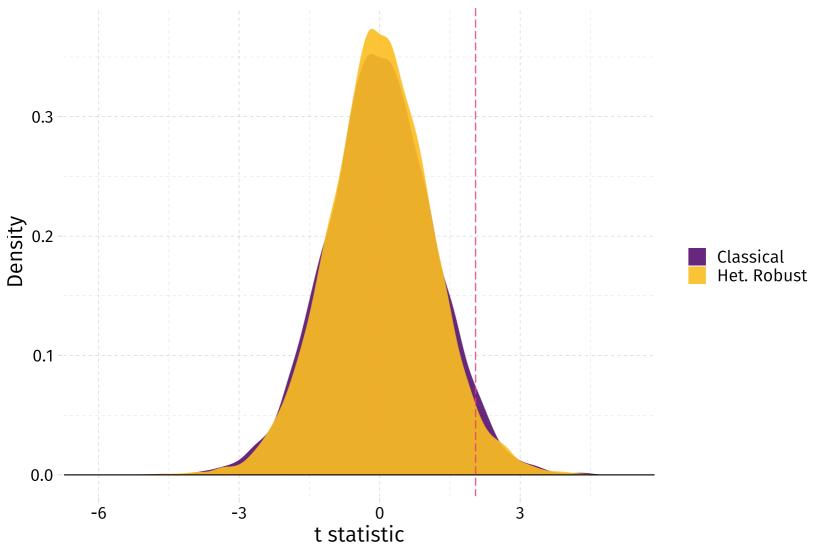
Now we run our new function flex\_iter() 10,000 times

```
# Packages
p load(purrr, furrr)
# Set options
set.seed(123)
# Tell R to parallelize
plan(multiprocess)
# Run 10,000 iterations
null df \leftarrow future map(
  1:1e4, flex iter,
  # Enforce the null hypothesis
  y = 0,
  # Specify heteroskedasticity
  \delta = 1,
  .options = future options(seed = T)
) %>% bind rows()
```

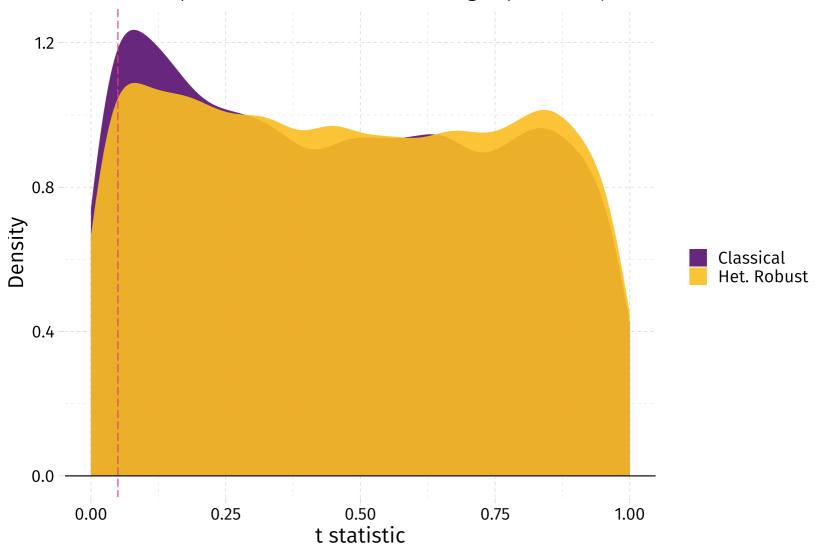
Comparing the distributions of standard errors for the coefficient on  $\boldsymbol{x}$ 



Comparing the distributions of t statistics for the coefficient on x



Distributions of p-values: both methods slightly over-reject the (true) null



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