Instrumental Variables

EC 421, Set 11

Edward Rubin

Prologue

Schedule

Last Time

Causality

Today

• Review: Causality

• New: Instrumental variables

Review

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$$\tau_i = y_{1i} - y_{0i}$$

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We cannot simultaneously know y_{1i} and y_{0i} .

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but never both at the same time.

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Idea: Estimate the **average treatment effect** as the difference between the average outcomes in the treatment group and the control group, *i.e.*,

$$Avg(y_i \mid D_i = 1) - Avg(y_i \mid D_i = 0)$$

where $D_i=1$ if i received treatment, and $D_i=0$ if i is in the control group.

Review

Result: We showed that even when the treatment effect is constant (meaning $\tau_i = \tau$ for all i),

$$egin{aligned} Avg(y_i \mid D_i = 1) - Avg(y_i \mid D_i = 0) \ &= au + \underbrace{Avg(y_{0,i} \mid D_i = 1) - Avg(y_{0,i} \mid D_i = 0)}_{ ext{Selection bias}} \end{aligned}$$

which says that the difference in the groups' means will give us a **biased estimate** for the causal effect of treatment **if we have selection bias.**

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A: (Formal) The average *untreated* outcome for a member of our **treatment group** (which we cannot observe) differs from the average *untreated* outcome for a member of our **control group**, *i.e.*,

$$Avg(y_{0,i} \mid D_i = 1) - Avg(y_{0,i} \mid D_i = 0)$$

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Practical problem: Selection bias is also difficult to observe

$$\underbrace{ rac{Avg(y_{0,i} \mid D_i = 1)}{ ext{Unobservable}} - Avg(y_{0,i} \mid D_i = 0) }_{ ext{Unobservable}}$$

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Bigger problem: If selection bias is present, our estimate for τ is biased, preventing us from understanding the causal effect of treatment.

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Bigger problem: If selection bias is present, our estimate for τ is biased, preventing us from understanding the causal effect of treatment.

Sounds a bit like omitted-variable bias, right? Our treatment variable is correlated with something that makese the two groups different.

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Example: Imagine we have two people—Al and Bri—and a single binary treatment, college. We interested in the effect of college on earnings.

$$Earn_{1,Al} = $60K$$

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but any real-world estimate would have serious selection issues since $Earn_{0,Al} \neq Earn_{0,Bri}$.

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- **Option 1: Distribute treatment** in a way such that the treatment and control groups are essentially identical (experiments).
- **Option 2: Build a control** group that *matches* the treatment group (life with observational data).

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- the **exogenous** part of x, which gives us unbiased estimates
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Instrumental variables attempts to separate out

- the **exogenous** part of x, which gives us unbiased estimates
- the **endogenous** part of x, which biases our results

If we use only the exogenous (good) variation in x, then we can avoid selection bias/omitted-variable bias.

Introductory example

Example: If we want to estimate the effect of veteran status on earnings,

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We would love to calculate $\frac{\text{Earnings}_{1i}}{\text{Earnings}_{0i}}$, but we can't.

And OLS will likely be biased for (1) due to selection/omitted-variable bias.

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$$\begin{aligned} \operatorname{Earnings}_{i} &= \beta_{0} + \beta_{1} \operatorname{Veteran}_{i} + u_{i} \\ &= \beta_{0} + \beta_{1} \left(\operatorname{Veteran}_{i}^{\operatorname{Exog.}} + \operatorname{Veteran}_{i}^{\operatorname{Endog.}} \right) + u_{i} \\ &= \beta_{0} + \beta_{1} \operatorname{Veteran}_{i}^{\operatorname{Exog.}} + \underbrace{\beta_{1} \operatorname{Veteran}_{i}^{\operatorname{Endog.}} + u_{i}}_{w_{i}} \\ &= \beta_{0} + \beta_{1} \operatorname{Veteran}_{i}^{\operatorname{Exog.}} + w_{i} \end{aligned}$$

$$(1)$$

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A₁: Choices to enlist in the military that are essentially random—or at least uncorrelated with omitted variables and the disturbance.

A₂: No selection bias:

$$Avg(\text{Earnings}_{0i} \mid \text{Veteran}_i = 1) - Avg(\text{Earnings}_{0i} \mid \text{Veteran}_i = 0) = 0$$

Instruments

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A: An **instrument** is a variable that is

- 1. correlated with the explanatory variable of interest (relevant),
- 2. uncorrelated with the disturbance (exogenous).

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So if we want an instrument z_i for endogenous veteran status in

$$\text{Earnings}_i = \beta_0 + \beta_1 \text{Veteran}_i + u_i$$

- 1. Relevant: $Cov(Veteran_i, z_i) \neq 0$
- 2. **Exogenous:** $Cov(z_i, u_i) = 0$

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- 3. Vietnam War draft

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the lottery was random

Instrumental review

Let's recap...

- Our instrument must be correlated with our endogenous variable.
- Our instrument must be uncorrelated with any other variable that affects the outcome.

Instrumental review

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- Our instrument must be correlated with our endogenous variable.
- Our instrument must be uncorrelated with any other variable that affects the outcome.

In other words:

The instrument only affects our outcome through the endogenous variable.

Back to our example

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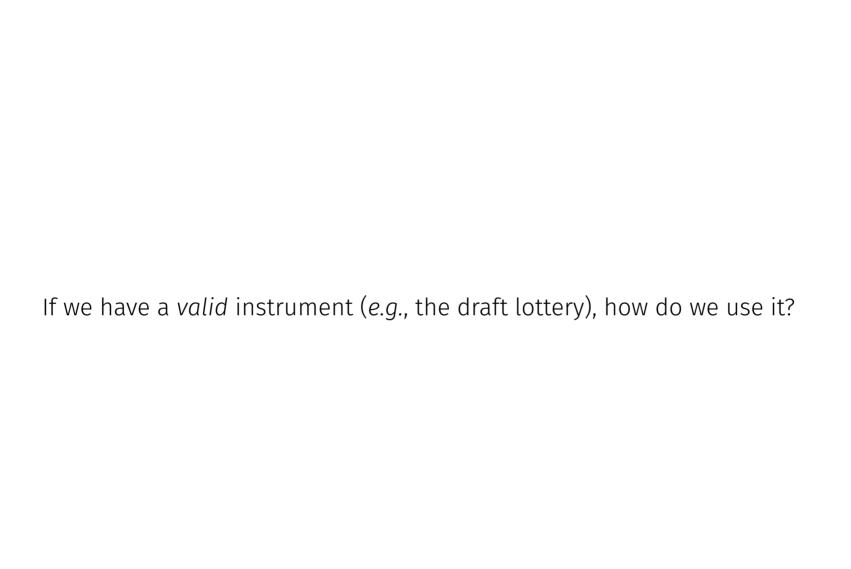
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Not exogenous

3. Vietnam War draft Relevant

Exogenous

Thus, only the Vietnam War's draft lottery appears to be a *valid* instrument.



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Let's consider two related effects:

1. The effect of the **instrument** on the **endogenous variable**, e.g.,

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1. The effect of the instrument on the endogenous variable, e.g.,

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2. The effect of the **instrument** on the **outcome variable**, e.g.,

$$\operatorname{Earnings}_i = \pi_0 + \pi_1 \operatorname{Draft}_i + w_i$$

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and we know that the draft affected veteran status.

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and we know that the draft affected veteran status.

Using our assumptions on independence and exogeneity:

```
(Effect of the draft on earnings) =
    (Effect of the draft on veteran status)*
    (Effect of veteran status on earnings)
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We just wrote out an expression for the effect of the draft on earnings, i.e.,

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```

but we want to know the effect of **veteran status** on **earnings**.

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```

but we want to know the effect of veteran status on earnings. Rearrange!

Estimation

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We just wrote out an expression for the effect of the draft on earnings, i.e.,
(Effect of the draft on earnings) =
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but we want to know the effect of veteran status on earnings. Rearrange!

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(Effect of the draft on earnings)
```

Our instrument consistently estimates both parts of this fraction!

(Effect of **the draft** on **veteran status**)

Estimation: Bring it all together

By estimating two regressions involving our **instrument**,

1. The effect of the **instrument** on the **endogenous variable**, e.g.,

$$Veteran_i = \gamma_0 + \gamma_1 Draft_i + v_i$$

2. The effect of the **instrument** on the **outcome variable**, e.g.,

$$\operatorname{Earnings}_i = \pi_0 + \pi_1 \operatorname{Draft}_i + w_i$$

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we can estimate our desired effect:

(Effect of **veteran status** on **earnings**) =
$$\frac{\pi_1}{\gamma_1}$$

Estimation: Bring it all together

So with instrumental variables, we estimate β_1 using

$${\hat eta}_1^{ ext{IV}} = rac{{\hat \pi}_1}{{\hat \gamma}_1}$$

where $\hat{\pi}_1$ and $\hat{\gamma}_1$ come from the two equations we just discussed.

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which equals β_1 as long as our instrument is **exogenous** (numerator) and **relevant** (denominator).

Figure 1

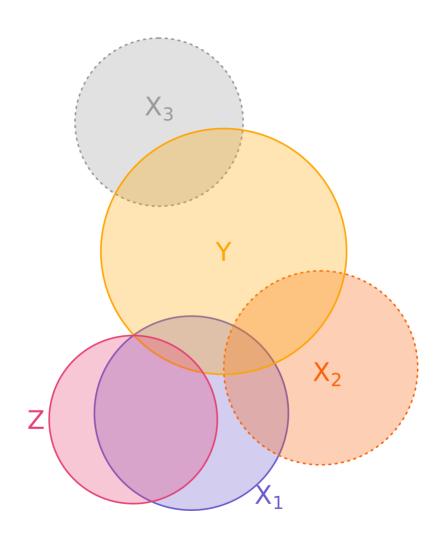
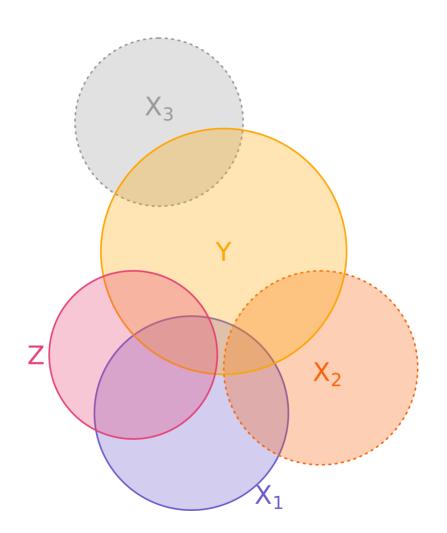


Figure 2



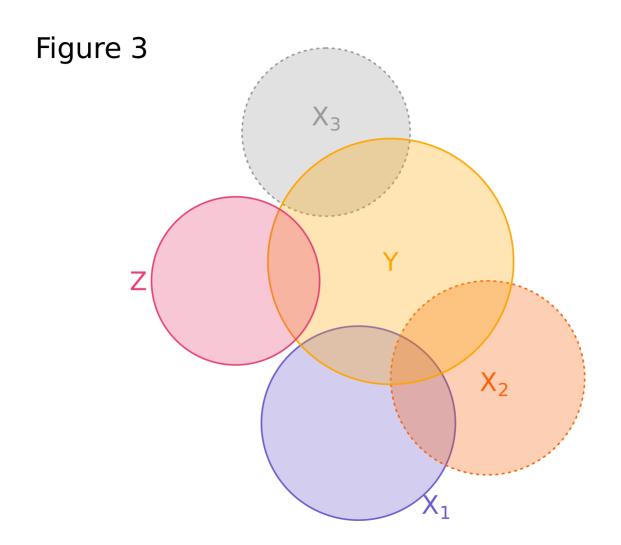
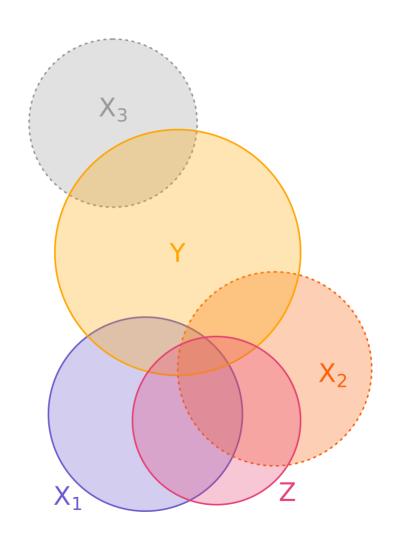


Figure 4



Venn diagram explanation

In these figures (Venn diagrams)

- Each circle illustrates a variable.
- Overlap gives the share of correlatation between two variables.
- Dotted borders denote omitted variables.

Take-aways

- Figure 1: Valid instrument (relevant; exogenous)
- Figure 2: **Invalid instrument** (relevant; not exogenous)
- Figure 3: **Invalid instrument** (not relevant; not exogenous)
- Figure 4: **Invalid instrument** (relevant; not exogenous)

Let's work an example in **R**.

Example in R

Back to our age-old battle to estimate the returns to education.

```
#> # A tibble: 722 × 4
#>
       wage education education_dad education_mom
#>
      <int>
                 <int>
                                <int>
                                                <int>
        769
                    12
                                     8
                                                    8
#>
#>
       808
                    18
                                    14
                                                   14
#>
       825
                    14
                                    14
                                                   14
#>
       650
                    12
                                    12
                                                   12
#>
       562
                    11
                                    11
                                                    6
                                                    8
#>
    6 600
                    10
                                     8
       1154
                                                   14
#>
                    15
#>
    8
       1000
                    12
                                    11
                                                   12
#>
       930
                    18
                                    14
                                                   13
#> 10
        900
                    15
                                    12
                                                   12
  # i 712 more rows
```

Example in R

OLS for the returns to education with will likely (definitely) be biased...

$$Wage_i = \beta_0 + \beta_1 Education_i + u_i$$

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but what if mother's education provides a valid instrument?

Example in R

We can check/test the relevance of mother's education for education.

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This regression is known as the *first stage*:

The effect of the instrument on our endogenous explanatory variable.

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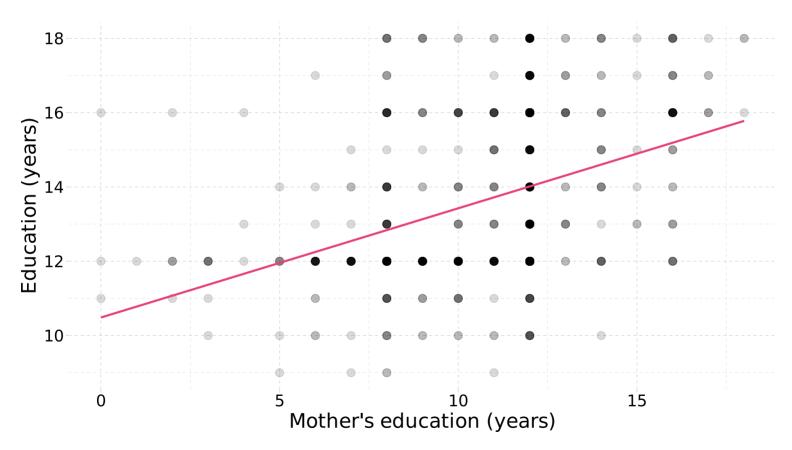
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The *p*-value suggests a very strong relationship (very *relevant*).

Visualizing the first stage



Visualizing the first stage



Exogeneity

Q: What does **exogeneity** mean in this case?

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A: We need two things

- 1. Mother's education (our instrument) must only affect earnings through education (our endogenous explanatory variable).
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Q: Does *mother's education* seem likely to satisfy exogeneity?

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Now, let's estimate the **reduced form**:

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Q₁: How do we interpret this estimated coefficient $(\hat{\pi}_1)$?

Q₂: If our instrument is *valid*, can we interpret these estimates as **causal**?

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- 1. In the **reduced-form equation**, we estimated $\hat{\pi}_1 \approx 31.81$.
- 2. In the **first-stage equation**, we estimated $\hat{\gamma}_1 \approx 0.294$.

$$\implies \hat{eta}_1^{ ext{IV}} = rac{\hat{\pi}_1}{\hat{\gamma}_1} = rac{31.81}{0.294} pprox 108.2$$

Example in R

Alternative: Use the function iv_robust() from the estimatr package.

This new function iv_robust works very similar to our good friend lm:

```
iv_robust(y \sim x \mid z, data = dataset)
```

- formula Specify the regression followed by | and your instrument (z).
- data You still need a dataset.

Note: As you might guess by its name, iv_robust calculates heteroskedasticity-robust standard errors by default.

Example in R

In practice...

```
# Estimate our IV regression
iv_est ← iv_robust(wage ~ education | education_mom, data = wage_df)
```

Term	Est.	S.E.	t stat.	p-Value
Intercept	-501.474	226.476	-2.21	0.0271
Education	108.214	16.810	6.44	

More

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A: Too bad. Extend IV to two-stage least squares (2SLS).

Intro

The intuition and insights of IV carry over into two-stage least squares.

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$$egin{aligned} ext{Endogenous model} & ext{Outcome}_i = eta_0 + eta_1 (ext{Endog. var.})_i + u_i \ & ext{First stage} & ext{(Endog. var.)}_i = \pi_0 + \pi_1 ext{Instrument}_i + v_i \ & ext{Second stage} & ext{Outcome}_i = \delta_0 + \delta_1 (ext{Endog. var.})_i + arepsilon_i \ & ext{Outcome}_i = \pi_0 + \pi_1 ext{Instrument}_i + w_i \end{aligned}$$

where $(Endog. var.)_i$ denotes the predicted values (*fitted values*) from the first-stage regression.

Intro

Two-stage least squares is very flexible—we include other controls, additional endogenous variables, and have multiple instruments.

But don't get too distracted by this fancy flexiblity.

We still need **valid** instruments.

In R

Back to our returns to education example.

$$Wage_i = \beta_0 + \beta_1 Education_i + u_i$$

Imagine that mother's and father's education are both valid instruments.

Then our **first-stage regression** is

$$Education_i = \gamma_0 + \gamma_1 (Mother's education)_i + \gamma_2 (Father's education)_i + v_i$$

which we can estimate via OLS.

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Q: Why?

In R

```
Education_i = \gamma_0 + \gamma_1 (Mother's education)_i + \gamma_2 (Father's education)_i + v_i
```

```
stage1 ← lm(education ~ education_mom + education_dad, wage_df)
```

First-stage results:

Term	Est.	S.E.	t stat.	p-Value
Intercept	9.845	0.305	32.31	
Mother's Education	0.149	0.032	4.62	
Father's Education	0.216	0.028	7.84	

In R

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\text{Education}_i = \gamma_0 + \gamma_1 (\text{Mother's education})_i + \gamma_2 (\text{Father's education})_i + v_i
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Our instruments each appear to be relevant.

Formally, we should jointly test them (e.g., F test).

In R

Using our estimated first stage, we grab the fitted endogenous variable

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```
# Add fitted values from first stage
wage_df$education_hat ← stage1$fitted.values
```

In R

Using our estimated first stage, we grab the fitted endogenous variable

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```
# Add fitted values from first stage
wage_df$education_hat ← stage1$fitted.values
```

Now we use OLS (again) to estimate the **second-stage regression**

$$\mathrm{Wage}_i = \delta_0 + \delta_1 \widehat{\mathrm{Education}}_i + \varepsilon_i$$

In R

$$Wage_i = \delta_0 + \delta_1 \widehat{Education}_i + \varepsilon_i$$

stage2 ← lm(wage ~ education_hat, wage_df)

Second-stage results:

Term	Est.	S.E.	t stat.	p-Value
Intercept	-454.683	198.149	-2.29	0.022
Fitted Education	104.789	14.462	7.25	

Ordinary least squares

Term	Est.	S.E.	t stat.	p-Value
Intercept	176.504	89.152	1.98	0.0481
Education	58.594	6.439	9.10	

Instrumental variables

Term	Est.	S.E.	t stat.	p-Value
Intercept	-501.474	226.476	-2.21	0.0271
Education	108.214	16.810	6.44	

Two-stage least squares w/ two instruments

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In R

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iv_robust(y \sim x1 + x2 + ... | z1 + z2 + ..., data)
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In our case, we have

- one explanatory variable (x) (education)
- two instruments (z) (parents' educations)

```
iv_robust(wage ~ education | education_mom + education_dad, data = wage_df)
```

Term	Est.	S.E.	t stat.	p-Value
Intercept	-454.683	199.946	-2.27	0.0233
Education, fitted	104.789	14.852	7.06	

There's more!

Because 2SLS **isolates exogenous variation in an endogenous variable**, we apply it in other settings that are biased from *endogenous* relationships.

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Common applications

- General causal inference for observational data (as we've seen).
- Experiments: Randomize a treatment that affects an endog. variable.
- ullet Measurement error: Regress noisy x_1 on noisy x_2 to capture signal.
- Simultaneous relationships (e.g., p and q from supply and demand).

However, in any 2SLS/IV setting, you need to mind the requirements for **valid instruments**—exogeneity and relevance.

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Instrumental variables

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- 6. Two-stage least squares
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 - Back to R
- 7. More applications