EC 421, Set 04

Edward Rubin Spring 2020

# Prologue

#### R showcase

#### R Markdown

- Simple mark-up language for for combining/creating documents, equations, figures, R, and more
- Basics of Markdown
- *E.g.*, \*\*I'm bold\*\*, \*I'm italic\*, I ← "code"

#### **Econometrics with** R

- (Currently) free, online textbook
- Written and published using R (and probably R Markdown)
- *Warning:* I haven't read this book yet.

Related: Tyler Ransom has a great cheatsheet for econometrics.

## Schedule

#### Last Time

We wrapped up our review.

#### Today

Heteroskedasticity

### Schedule

#### This week

First assignment! Due Sunday-don't wait.

Turn in **2 files**<sup>†</sup>

- 1. Your write up (*e.g.*, Word file).
- 2. The R script that generated your answers.

#### Important

- We should be able to easily find your answers for each question.
- **Do not copy.** (You will receive a zero.)

t: Unless you're using RMarkdown—then we need a PDF or HTML file.

## Schedule

#### The future

- Next assignment: Next week
- Midterm: In two weeks

Let's write down our **current assumptions** 

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6. The disturbances come from a **Normal** distribution, *i.e.*,  $u_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$ .

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#### Violation of this assumption:

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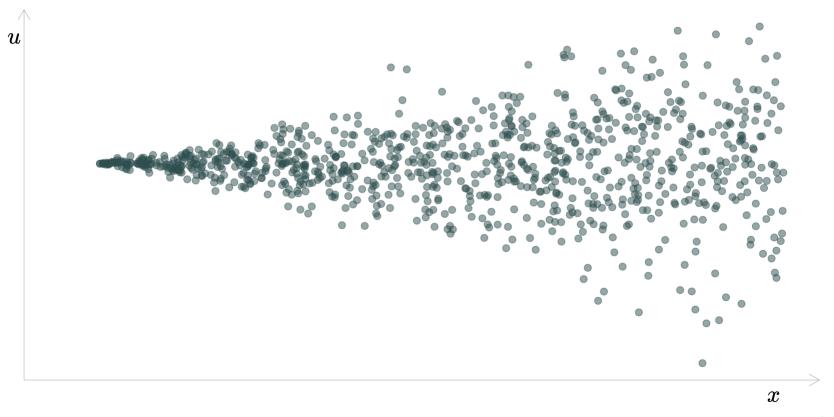
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In other words: Our disturbances have different variances.

Classic example of heteroskedasticity: The funnel

Variance of u increases with x



Another example of heteroskedasticity: (double funnel?)

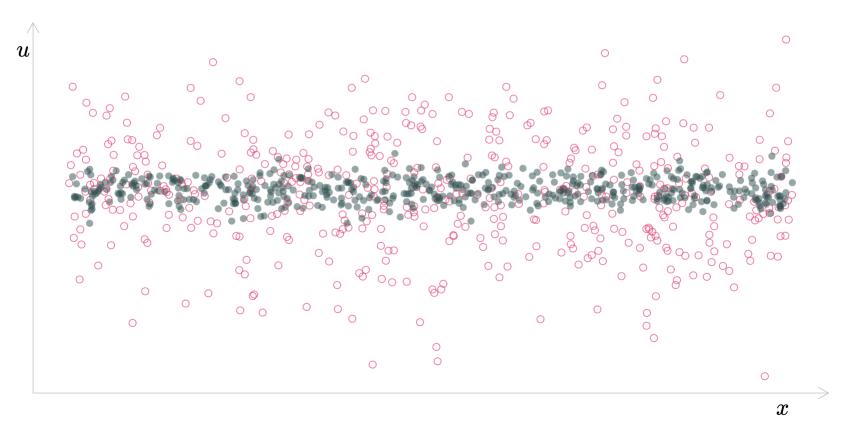
Variance of u increasing at the extremes of x



x

Another example of heteroskedasticity:

Differing variances of u by group



**Heteroskedasticity** is present when the variance of u changes with any combination of our explanatory variables  $x_1$ , through  $x_k$  (henceforth: X).

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**Why we care:** Heteroskedasticity shows us how small violations of our assumptions can affect OLS's performance.

#### Consequences

So what are the consquences of heteroskedasticity? Bias? Inefficiency?

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**Recall<sub>2</sub>:** We previously showed 
$$\hat{eta}_1 = rac{\sum_i \left(y_i - \overline{y}
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It will actually help us to rewrite this estimator as

$${\hat eta}_1 = eta_1 + rac{{\sum _i \left( {{x_i} - \overline x} 
ight){u_i}}}{{\sum _i \left( {{x_i} - \overline x} 
ight)^2 }}$$

**Proof:** Assuming  $y_i = eta_0 + eta_1 x_i + u_i$ 

$$egin{aligned} \hat{eta}_1 &= rac{\sum_i \left(y_i - ar{y}
ight) \left(x_i - ar{x}
ight)^2}{\sum_i \left(x_i - ar{x}
ight)^2} \ &= rac{\sum_i \left(\left[eta_0 + eta_1 x_i + u_i
ight] - \left[eta_0 + eta_1 ar{x} + ar{u}
ight]
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ight] \left[u_i - ar{u}
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ight)}{\sum_i \left(x_i - ar{x}
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$$\begin{split} \hat{\beta}_{1} &= \dots = \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) \left(u_{i} - \overline{u}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \sum_{i} \left(x_{i} - \overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - \sum_{i} \overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - n\overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - \sum_{i} x_{i}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i}}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \end{split}$$

#### Consequences: Bias

We now want to see if heteroskedasticity biases the OLS estimator for  $\beta_1$ .

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Phew. **OLS is still unbiased** for the  $\beta_k$ .

#### **Consequences: Efficiency**

OLS's efficiency and inference do not survive heteroskedasticity.

• In the presence of heteroskedasticity, OLS is **no longer the most efficient** (best) linear unbiased estimator.

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- In the presence of heteroskedasticity, OLS is **no longer the most efficient** (best) linear unbiased estimator.
- It would be more informative (efficient) to weight observations inversely to their  $u_i$ 's variance.
  - Downweight high-variance  $u_i$ 's (too noisy to learn much).
  - Upweight observations with low-variance  $u_i$ 's (more 'trustworthy').
  - Now you have the idea of weighted least squares (WLS)

#### **Consequences: Inference**

OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)

## Heteroskedasticity

#### **Consequences: Inference**

OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)
- It's hard to learn much without sound inference.

## Heteroskedasticity

#### Solutions

- 1. **Tests** to determine whether heteroskedasticity is present.
- 2. **Remedies** for (1) efficiency and (2) inference

While we *might* have solutions for heteroskedasticity, the efficiency of our estimators depends upon whether or not heteroskedasticity is present.

- 1. The Goldfeld-Quandt test
- 2. The Breusch-Pagan test
- 3. The White test

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Each of these tests centers on the fact that we can **use the OLS residual**  $e_i$ to estimate the population disturbance  $u_i$ .

#### The Goldfeld-Quandt test

Focuses on a specific type of heteroskedasticity: whether the variance of  $u_i$  differs **between two groups**.<sup>†</sup>

Remember how we used our residuals to estimate the  $\sigma^2$ ?

$$s^2 = rac{ ext{SSE}}{n-1} = rac{\sum_i e_i^2}{n-1}$$

We will use this same idea to determine whether there is evidence that our two groups differ in the variances of their disturbances, effectively comparing  $s_1^2$  and  $s_2^2$  from our two groups.

[+]: The G-Q test was one of the early tests of heteroskedasticity (1965).

#### The Goldfeld-Quandt test

Operationally,

- 1. Order your the observations by  $\boldsymbol{x}$
- 2. Split the data into two groups of size  $\texttt{n}^{\star}$ 
  - $\circ$  G\_1: The first third
  - $\circ$  G\_2: The last third
- 3. Run separate regressions of y on x for  $\mathsf{G}_1$  and  $\mathsf{G}_2$
- 4. Record  $\ensuremath{\mathsf{SSE}}_1$  and  $\ensuremath{\mathsf{SSE}}_2$
- 5. Calculate the G-Q test statistic

#### The Goldfeld-Quandt test

The G-Q test statistic

$$F_{(n^{\star}-k,\,n^{\star}-k)} = rac{\mathrm{SSE}_2/(n^{\star}-k)}{\mathrm{SSE}_1/(n^{\star}-k)} = rac{\mathrm{SSE}_2}{\mathrm{SSE}_1}$$

follows an F distribution (under the null hypothesis) with  $n^{\star} - k$  and  $n^{\star} - k$  degrees of freedom.<sup>†</sup>

#### The Goldfeld-Quandt test

The G-Q test statistic

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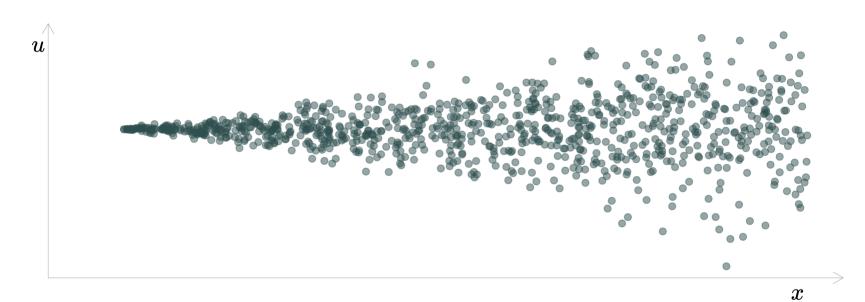
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#### Notes

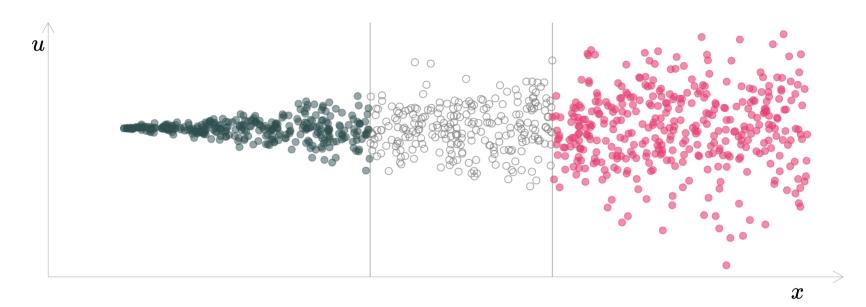
- The G-Q test requires the disturbances follow normal distributions.
- The G-Q assumes a very specific type/form of heteroskedasticity.
- Performs very well if we know the form of potentially heteroskedasticity.

[+]: Goldfeld and Quandt suggested  $n^*$  of (3/8)n. k gives number of estimated parameters (*i.e.*,  $\hat{\beta}_i$ 's).

#### The Goldfeld-Quandt test



#### The Goldfeld-Quandt test



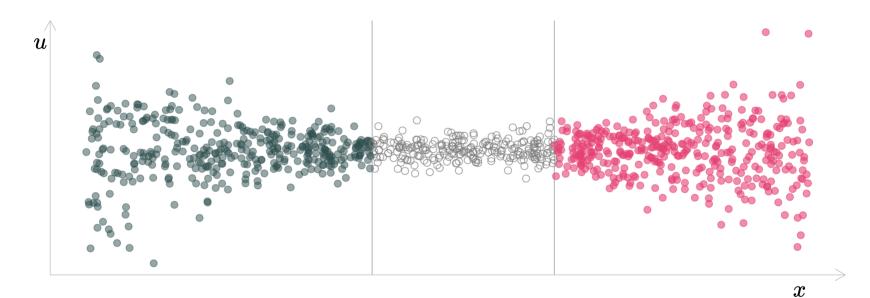
$$F_{375,\,375} = rac{{
m SSE}_2 = 18,\,203.4}{{
m SSE}_1 = 1,\,039.5} pprox 17.5 \implies p ext{-value} < 0.001$$

: We reject  $H_0$ :  $\sigma_1^2 = \sigma_2^2$  and conclude there is statistically significant evidence of heteroskedasticity.

#### The Goldfeld-Quandt test

The problem...

#### The Goldfeld-Quandt test



$$F_{375,\,375} = rac{\mathrm{SSE}_2 = 14,516.8}{\mathrm{SSE}_1 = 14,937.1} pprox 1 \implies p$$
-value  $pprox 0.609$ 

 $\therefore$  We fail to reject H<sub>0</sub>:  $\sigma_1^2 = \sigma_2^2$  while heteroskedasticity is present.

#### The Breusch-Pagan test

Breusch and Pagan (1981) attempted to solve this issue of being too specific with the functional form of the heteroskedasticity.

- Allows the data to show if/how the variance of  $u_i$  correlates with X.
- If  $\sigma_i^2$  correlates with X, then we have heteroskedasticity.
- Regresses  $e_i^2$  on  $X = [1, x_1, x_2, \ldots, x_k]$  and tests for joint significance.

#### The Breusch-Pagan test

How to implement:

1. Regress y on an intercept, 
$$x_1$$
,  $x_2$ , ...,  $x_k$ .

- 2. Record residuals e.
- 3. Regress  $e^2$  on an intercept,  $x_1$ ,  $x_2$ , ...,  $x_k$ .

$$e_i^2=lpha_0+lpha_1x_{1i}+lpha_2x_{2i}+\dots+lpha_kx_{ki}+v_i$$

4. Record  $R^2$ .

5. Test hypothesis H $_0$ :  $lpha_1=lpha_2=\dots=lpha_k=0$ 

#### The Breusch-Pagan test

The B-P test statistic<sup>†</sup> is

$${
m LM}=n imes R_e^2$$

where  $R_e^2$  is the  $R^2$  from the regression

$$e_i^2=lpha_0+lpha_1x_{1i}+lpha_2x_{2i}+\dots+lpha_kx_{ki}+v_i$$

Under the null, LM is asymptotically distributed as  $\chi^2_k$ .

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This test statistic tests  $\mathsf{H}_0: \alpha_1 = \alpha_2 = \cdots = \alpha_k = 0$ .

Rejecting the null hypothesis implies evidence of heteroskedasticity.

[+]: This specific form of the test statistic actually comes form Koenker (1981).

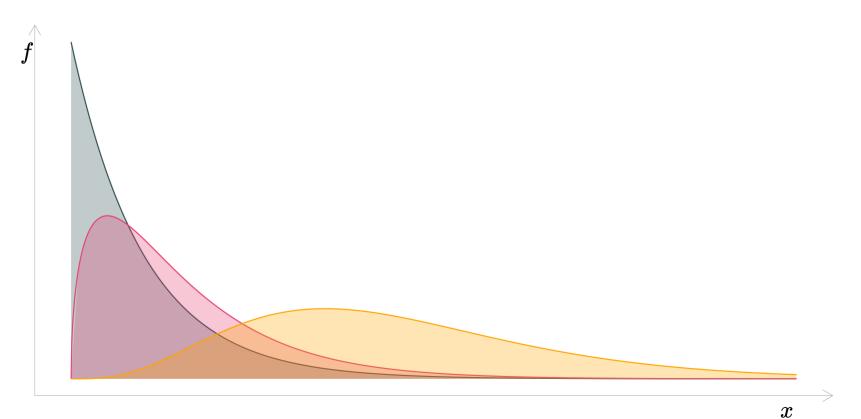
#### The $\chi^2$ distribution

We just mentioned that under the null, the B-P test statistic is distributed as a  $\chi^2$  random variable with k degrees of freedom.

The  $\chi^2$  distribution is just another example of a common (named) distribution (like the Normal distribution, the *t* distribution, and the *F*).

#### The $\chi^2$ distribution

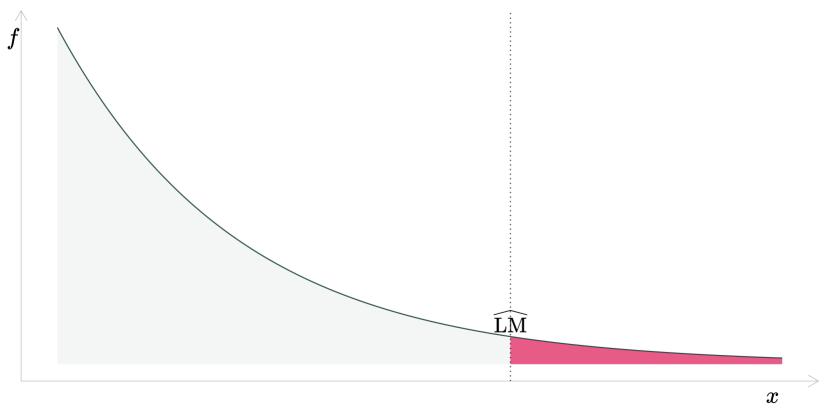
Three examples of  $\chi^2_k$ : k=1, k=2, and k=9



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#### The $\chi^2$ distribution

Probability of observing a more extreme test statistic  $\widehat{\mathbf{LM}}$  under  $\mathsf{H}_0$ 



#### The Breusch-Pagan test

**Problem:** We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances  $\sigma_i^2$ .

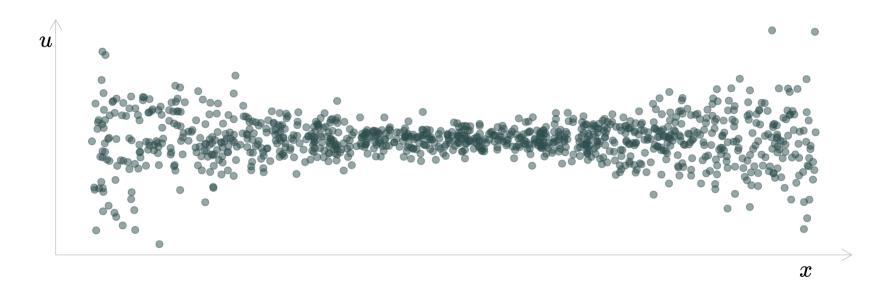
#### The Breusch-Pagan test

**Problem:** We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances  $\sigma_i^2$ .

**Result:** B-P *may* still miss fairly simple forms of heteroskedasticity.

#### The Breusch-Pagan test

Breusch-Pagan tests are still **sensitive to functional form**.



$$egin{aligned} e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} & \widehat{ ext{LM}} &= 1.26 & p ext{-value} &pprox 0.261 \ e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} + \hat{lpha}_2 x_{1i}^2 & \widehat{ ext{LM}} &= 185.8 & p ext{-value} &< 0.001 \end{aligned}$$

#### The White test

So far we've been testing for specific relationships between our explanatory variables and the variances of the disturbances, *e.g.*,

- $\mathsf{H}_0: \sigma_1^2 = \sigma_2^2$  for two groups based upon  $x_j$  (**G-Q**)
- $\mathsf{H}_0: \alpha_1 = \cdots = \alpha_k = 0$  from  $e_i^2 = \alpha_0 + \alpha_1 x_{1i} + \cdots + \alpha_k x_{ki} + v_i$  (**B-P**)

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However, we actually want to know if

$$\sigma_1^2=\sigma_2^2=\dots=\sigma_n^2$$

**Q:** Can't we just test this hypothesis? **A:** Sort of.

#### The White test

Toward this goal, Hal White took advantage of the fact that we can **replace the homoskedasticity requirement with a weaker assumption**:

- Old:  $\operatorname{Var}(u_i|X) = \sigma^2$
- New:  $u^2$  is uncorrelated with the explanatory variables (*i.e.*,  $x_j$  for all j), their squares (*i.e.*,  $x_j^2$ ), and the first-degree interactions (*i.e.*,  $x_jx_h$ ).

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Toward this goal, Hal White took advantage of the fact that we can **replace the homoskedasticity requirement with a weaker assumption**:

- Old:  $\operatorname{Var}(u_i|X) = \sigma^2$
- New:  $u^2$  is uncorrelated with the explanatory variables (*i.e.*,  $x_j$  for all j), their squares (*i.e.*,  $x_j^2$ ), and the first-degree interactions (*i.e.*,  $x_jx_h$ ).

This new assumption is easier to explicitly test (*hint*: regression).

#### The White test

An outline of White's test for heteroskedasticity:

1. Regress y on  $x_1$ ,  $x_2$ , ...,  $x_k$ . Save residuals e.

2. Regress squared residuals on all explanatory variables, their squares, and interactions.

$$e^2 = lpha_0 + \sum_{h=1}^k lpha_h x_h + \sum_{j=1}^k lpha_{k+j} x_j^2 + \sum_{\ell=1}^{k-1} \sum_{m=\ell+1}^k lpha_{\ell,m} x_\ell x_m + v_i$$

3. Record  $R_e^2$ .

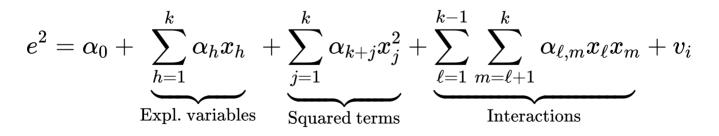
4. Calculate test statistic to test  $extsf{H}_0: \; lpha_p = 0$  for all p 
eq 0 .

#### The White test

Just as with the Breusch-Pagan test, White's test statistic is

$$\mathrm{LM} = n imes R_e^2 \qquad \mathrm{Under} \ \mathrm{H}_0, \ \mathrm{LM} \stackrel{\mathrm{d}}{\sim} \chi_k^2$$

but now the  $R_e^2$  comes from the regression of  $e^2$  on the explanatory variables, their squares, and their interactions.



**Note:** The k (for our  $\chi_k^2$ ) equals the number of estimated parameters in the regression above (the  $\alpha_j$ ), excluding the intercept ( $\alpha_0$ ).

#### The White test

**Practical note:** If a variable is equal to its square (*e.g.*, binary variables), then you don't (can't) include it. The same rule applies for interactions.

#### The White test

*Example:* Consider the model  $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u$ 

**Step 1:** Estimate the model; obtain residuals (e).

**Step 2:** Regress  $e^2$  on explanatory variables, squares, and interactions.

$$e^2 = lpha_0 + lpha_1 x_1 + lpha_2 x_2 + lpha_3 x_3 + lpha_4 x_1^2 + lpha_5 x_2^2 + lpha_6 x_3^2 \ + lpha_7 x_1 x_2 + lpha_8 x_1 x_3 + lpha_9 x_2 x_3 + v$$

Record the  $R^2$  from this equation (call it  $R_e^2$ ).

**Step 3:** Test  $\mathsf{H}_0$ :  $lpha_1 = lpha_2 = \dots = lpha_9 = 0$  using  $\mathrm{LM} = n R_e^2 \stackrel{\mathrm{d}}{\sim} \chi_9^2$ .

[†]: To simplify notation here, I'm dropping the *i* subscripts.

#### The White test



We've already done the White test for this simple linear regression.

$$e_i^2 = \hatlpha_0 + \hatlpha_1 x_{1i} + \hatlpha_2 x_{1i}^2 \qquad \widehat{ ext{LM}} = 185.8 \qquad p ext{-value} < 0.001$$

# Testing for Heteroskedasticity Examples

#### Examples

**Goal:** Estimate the relationship between standardized test scores (outcome variable) and (1) student-teacher ratio and (2) income, *i.e.*,

$$(\text{Test score})_i = \beta_0 + \beta_1 \text{Ratio}_i + \beta_2 \text{Income}_i + u_i$$
 (1)

**Potential issue:** Heteroskedasticity... and we do not observe  $u_i$ .

#### Solution:

- 1. Estimate the relationship in (1) using OLS
- 2. Use the residuals  $(e_i)$  to test for heteroskedasticity
  - Goldfeld-Quandt
  - Breusch-Pagan
  - White

#### Examples

We will use testing data from the dataset Caschool in the Ecdat R package.

### Examples

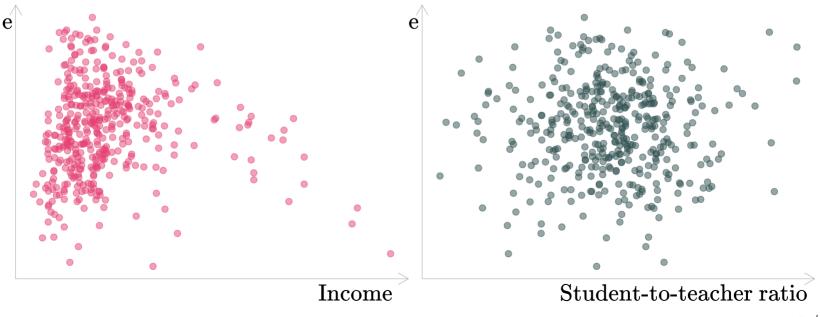
Let's begin by estimating our model

```
(	ext{Test score})_i = eta_0 + eta_1	ext{Ratio}_i + eta_2	ext{Income}_i + u_i
```

| #> | # | A tibble: 3 | x 5         |             |             |             |
|----|---|-------------|-------------|-------------|-------------|-------------|
| #> |   | term        | estimate    | std.error   | statistic   | p.value     |
| #> |   | <chr></chr> | <dbl></dbl> | <dbl></dbl> | <dbl></dbl> | <dbl></dbl> |
| #> | 1 | (Intercept) | 639.        | 7.45        | 85.7        | 5.70e-267   |
| #> | 2 | ratio       | -0.649      | 0.354       | -1.83       | 6.79e- 2    |
| #> | 3 | income      | 1.84        | 0.0928      | 19.8        | 4.38e- 62   |

### Examples

Now, let's see what the residuals suggest about heteroskedasticity



### Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

# Arrange the data by income
test\_df ← arrange(test\_df, income)

### Example: Goldfeld-Quandt

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# Arrange the data by income
test\_df ← arrange(test\_df, income)
# Re-estimate the model for the last and first 158 observations
est\_model1 ← lm(test\_score ~ ratio + income, data = tail(test\_df, 158))
est\_model2 ← lm(test\_score ~ ratio + income, data = head(test\_df, 158))

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e\_model1 ← residuals(est\_model1)
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# Grab the residuals from each regression
e_model1 ← residuals(est_model1)
e_model2 ← residuals(est_model2)
# Calculate SSE for each regression
(sse model1 ← sum(e model1^2))
```

#> [1] 19305.01

```
(sse_model2 ← sum(e_model2<sup>2</sup>))
```

### Example: Goldfeld-Quandt

Remember the Goldfeld-Quandt test statistic?

$$F_{n^{\star}-k,\,n^{\star}-k}=rac{\mathrm{SSE}_2}{\mathrm{SSE}_1}$$

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(f\_gq ← sse\_model2/sse\_model1)

#> [1] 1.530061

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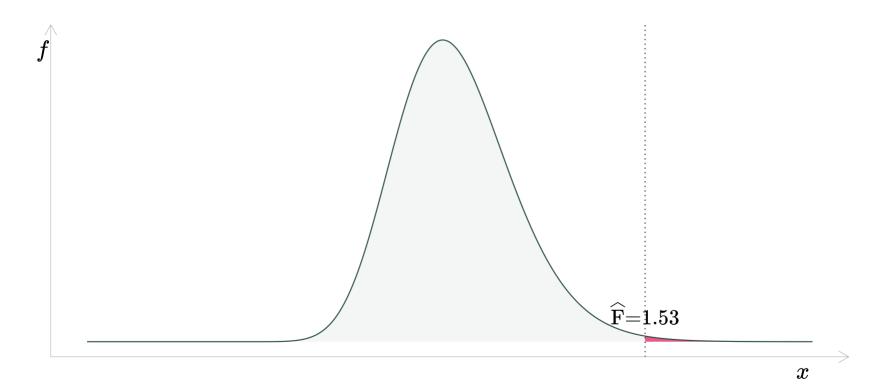
#> [1] 1.530061

# p-value
pf(q = f\_gq, df1 = 158-3, df2 = 158-3, lower.tail = F)

#> [1] 0.004226666

### Example: Goldfeld-Quandt

The Goldfeld-Quandt test statistic and its null distribution



### Example: Goldfeld-Quandt

Putting it all together:

 $extsf{H}_{0}\!\!:\sigma_{1}^{2}=\sigma_{2}^{2}$  vs.  $extsf{H}_{ extsf{A}}\!\!:\sigma_{1}^{2}
eq\sigma_{2}^{2}$ 

Goldfeld-Quandt test statistic: Fpprox 1.53

 $p ext{-value} pprox 0.00423$ 

 $\therefore$  Reject H<sub>0</sub> (*p*-value is less than 0.05).

**Conclusion:** There is statistically significant evidence that  $\sigma_1^2 \neq \sigma_2^2$ . Therefore, we find statistically significant evidence of heteroskedasticity (at the 5-percent level).

### Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

### Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

```
# Arrange the data by ratio
test_df ← arrange(test_df, ratio)
# Re-estimate the model for the last and first 158 observations
est_model3 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model4 ← lm(test_score ~ ratio + income, data = head(test_df, 158))
# Grab the residuals from each regression
e_model3 ← residuals(est_model3)
e_model4 ← residuals(est_model4)
# Calculate SSE for each regression
(sse model3 ← sum(e model3^2))
```

#> [1] 26243.52

```
(sse_model4 ← sum(e_model4<sup>2</sup>))
```

### Example: Goldfeld-Quandt

 $F_{n^{\star}-k,\,n^{\star}-k} = rac{\mathrm{SSE}_4}{\mathrm{SSE}_3} pprox rac{29,101.52}{26,243.52} pprox 1.11$ 

which has a *p*-value of approximately 0.2603.

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**Lesson:** Understand the limitations of estimators, tests, *etc.* 

#### Example: Breusch-Pagan

Let's test the same model with the Breusch Pagan.

Recall: We saved our residuals as e in our dataset, i.e.,

test\_df\$e ← residuals(est\_model)

### Example: Breusch-Pagan

In B-P, we first regress  $e_i^2$  on the explanatory variables,

### Example: Breusch-Pagan

and use the resulting  $R^2$  to calculate a test statistic.

```
# Regress squared residuals on explanatory variables
bp_model ← lm(I(e^2) ~ ratio + income, data = test_df)
# Grab the R-squared
(bp_r2 ← summary(bp_model)$r.squared)
```

#> [1] 3.23205e-05

#### Example: Breusch-Pagan

The Breusch-Pagan test statistic is

 ${
m LM}=n imes R_e^2$ 

#### Example: Breusch-Pagan

- The Breusch-Pagan test statistic is
- ${
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```
The Breusch-Pagan test statistic is
```

 ${
m LM}=n imes R_e^2pprox 420 imes 0.0000323pprox 0.0136$ 

which we test against a  $\chi^2_k$  distribution (here: k=2).<sup>†</sup>

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```
The Breusch-Pagan test statistic is
```

 ${
m LM}=n imes R_e^2pprox 420 imes 0.0000323pprox 0.0136$ 

which we test against a  $\chi^2_k$  distribution (here: k=2).<sup>+</sup>

#> [1] 0.9932357

[+]: k is the number of explanatory variables (excluding the intercept).

#### Example: Breusch-Pagan

 $\mathsf{H}_{0}\!\!:lpha_{1}=lpha_{2}=0$  vs.  $\mathsf{H}_{\mathsf{A}}\!\!:lpha_{1}
eq 0$  and/or  $lpha_{2}
eq 0$ 

for the model  $u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$ 

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p-value pprox 0.993

 $\therefore$  Fail to reject H<sub>0</sub> (the *p*-value is greater than 0.05)

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**Conclusion:** We do not find statistically significant evidence of heteroskedasticity at the 5-percent level.

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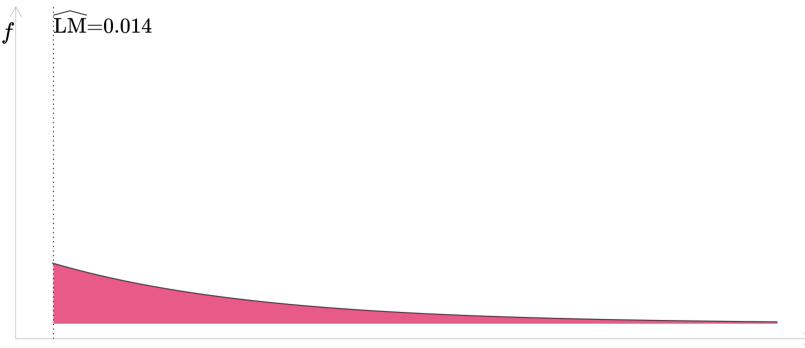
 $p ext{-value} pprox 0.993$ 

 $\therefore$  Fail to reject H<sub>0</sub> (the *p*-value is greater than 0.05)

**Conclusion:** We do not find statistically significant evidence of heteroskedasticity at the 5-percent level. (We find no evidence of a *linear* relationship between  $u_i^2$  and the explanatory variables.)

### Example: Breusch-Pagan

The Breusch-Pagan test statistic and its null distribution



### Heteroskedasticity

### Example: White

The White test adds squared terms and interactions to the B-P test.

$$egin{aligned} &u_i^2 =& lpha_0 + lpha_1 ext{Ratio}_i + lpha_2 ext{Income}_i \ &+ lpha_3 ext{Ratio}_i^2 + lpha_4 ext{Income}_i^2 + lpha_5 ext{Ratio}_i imes ext{Income}_i \ &+ w_i \end{aligned}$$

which moves the null hypothesis from H<sub>0</sub>:  $\alpha_1 = \alpha_2 = 0$  to H<sub>0</sub>:  $\alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = \alpha_5 = 0$ 

### Heteroskedasticity

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which moves the null hypothesis from H<sub>0</sub>:  $\alpha_1 = \alpha_2 = 0$  to H<sub>0</sub>:  $\alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = \alpha_5 = 0$ 

So we just need to update our R code, and we're set.

### Heteroskedasticity

### Example: White

Aside: R has funky notation for squared terms and interactions in lm():

- **Squared terms** use I(), *e.g.*,  $lm(y ~ I(x^2))$
- Interactions use : between the variables, *e.g.*, lm(y ~ x1:x2)

*Example*: Regress y on quadratic of x1 and x2:

```
# Pretend quadratic regression w/ interactions
lm(y ~ x1 + x2 + I(x1<sup>2</sup>) + I(x2<sup>2</sup>) + x1:x2, data = pretend_df)
```

#### Example: White

#### **Step 1:** Regress $e_i^2$ on 1<sup>st</sup> degree, 2<sup>nd</sup> degree, and interactions

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

#### Example: White

#### **Step 2:** Collect $R_e^2$ from the regression.

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

#> [1] 0.07332222

#### Example: White

#### **Step 3:** Calculate White test statistic $\mathrm{LM} = n imes R_e^2 pprox 420 imes 0.073$

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
white_r2 ← summary(white_model)$r.squared
# Calculate the White test statistic
(white stat ← 420 * white r2)
```

#> [1] 30.79533

#### Example: White

#### **Step 4:** Calculate the associated *p*-value (where $\operatorname{LM} \stackrel{d}{\sim} \chi^2_k$ ); here, k=5

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
white_r2 ← summary(white_model)$r.squared
# Calculate the White test statistic
white_stat ← 420 * white_r2
# Calculate the p-value
pchisq(q = white_stat, df = 5, lower.tail = F)
```

#> [1] 1.028039e-05

#### Example: White

Putting everything together...

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 $oxdota_0: lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ 

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Our White test statistic:  ${
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Under the  $\chi^2_5$  distribution, this  $\widehat{\mathrm{LM}}$  has a *p*-value less than 0.001.

#### Example: White

Putting everything together...

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: We reject H<sub>0</sub>

#### Example: White

Putting everything together...

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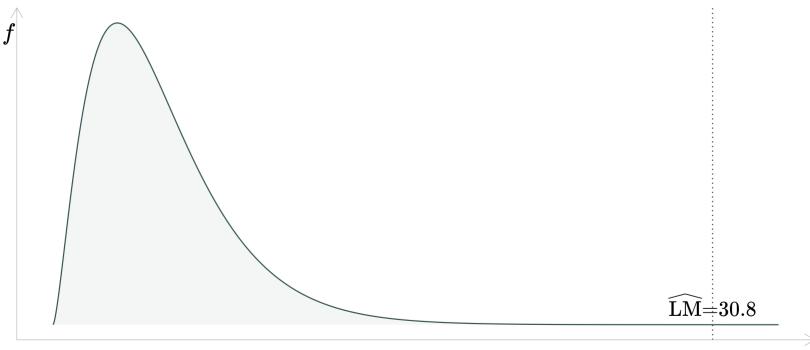
Our White test statistic:  ${
m LM}=n imes R_e^2pprox 420 imes 0.073pprox 30.8$ 

Under the  $\chi^2_5$  distribution, this  $\widehat{\mathrm{LM}}$  has a *p*-value less than 0.001.

∴ We **reject H**<sup>0</sup> and conclude there is **statistically significant evidence of heteroskedasticity** (at the 5-percent level).

#### Example: White

The White test statistic and its null distribution



- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting *y* against *x*, tell us anything about heteroskedasticity?
- **Q:** Does plotting *e* against *x*, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the  $u_i$ 's, what do we use to *learn about* heteroskedasticity?
- **Q:** Which test do you recommend to test for heteroskedasticity? Why?

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- A:

Math:  $\operatorname{Var}(u_i|X) \neq \operatorname{Var}(u_j|X)$  for some  $i \neq j$ .

**Words:** There is a systematic relationship between the variance of  $u_i$  and our explanatory variables.

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- A: It biases our standard errors—wrecking our statistical tests and confidence intervals. Also: OLS is no longer the most efficient (best) linear unbiased estimator.

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- A: It's not exactly what we want, but since y is a function of x and u, it can still be informative. If y becomes more/less disperse as x changes, we likely have heteroskedasticity.

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- **A:** Yes. The spread of *e* depicts its variance—and tells us something about the variance of *u*. Trends in this variance, along *x*, suggest heteroskedasticity.

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- A: We use the  $e_i$ 's to predict/learn about the  $u_i$ 's. This trick is key for almost everything we do with heteroskedasticity testing/correction.

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- **Q:** Which test do you recommend to test for heteroskedasticity? Why?
- A: I like White. Fewer assumptions. Fewer issues.