EC 421, Set 04

Edward Rubin Spring 2020

Prologue

R showcase

R Markdown

- Simple mark-up language for for combining/creating documents, equations, figures, R, and more
- Basics of Markdown
- *E.g.*, **I'm bold**, *I'm italic*, I ← "code"

Econometrics with R

- (Currently) free, online textbook
- Written and published using R (and probably R Markdown)
- *Warning:* I haven't read this book yet.

Related: Tyler Ransom has a great cheatsheet for econometrics.

Schedule

Last Time

We wrapped up our review.

Today

Heteroskedasticity

Schedule

This week

First assignment! Due Sunday-don't wait.

Turn in **2 files**[†]

- 1. Your write up (*e.g.*, Word file).
- 2. The R script that generated your answers.

Important

- We should be able to easily find your answers for each question.
- **Do not copy.** (You will receive a zero.)

t: Unless you're using RMarkdown—then we need a PDF or HTML file.

Schedule

The future

- Next assignment: Next week
- Midterm: In two weeks

Let's write down our **current assumptions**

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6. The disturbances come from a **Normal** distribution, *i.e.*, $u_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$.

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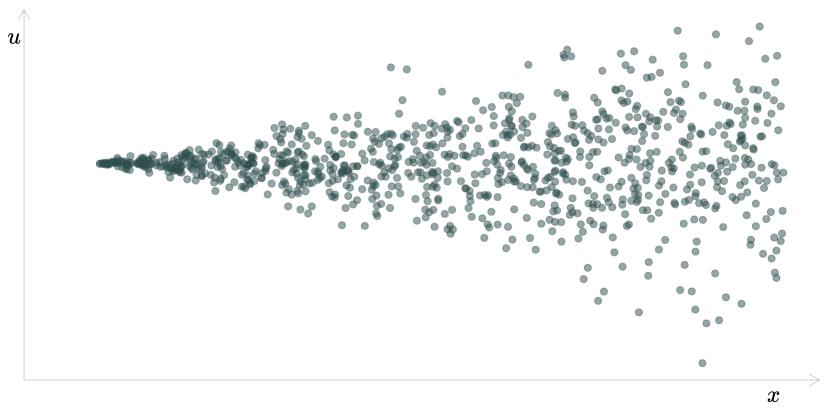
Violation of this assumption:

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$$\mathrm{Var}(u_i)=\sigma_i^2$$
 and $\sigma_i^2
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eq j.$

In other words: Our disturbances have different variances.

Classic example of heteroskedasticity: The funnel

Variance of u increases with x



Another example of heteroskedasticity: (double funnel?)

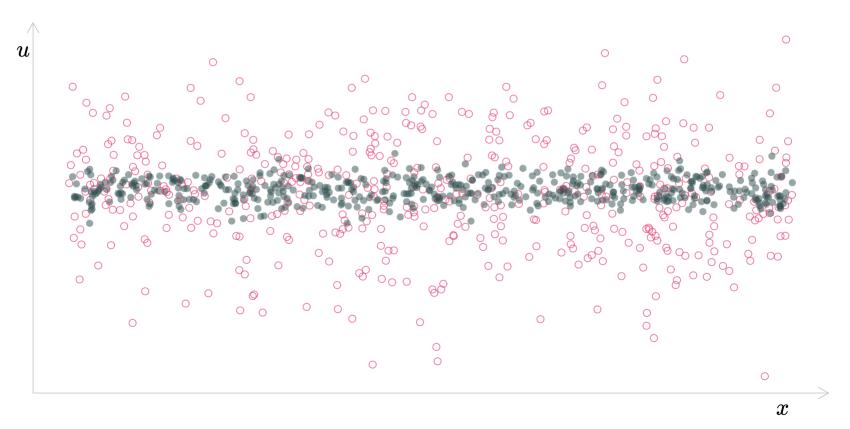
Variance of u increasing at the extremes of x



x

Another example of heteroskedasticity:

Differing variances of u by group



Heteroskedasticity is present when the variance of u changes with any combination of our explanatory variables x_1 , through x_k (henceforth: X).

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Why we care: Heteroskedasticity shows us how small violations of our assumptions can affect OLS's performance.

Consequences

So what are the consquences of heteroskedasticity? Bias? Inefficiency?

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Recall₂: We previously showed
$$\hat{eta}_1 = rac{\sum_i \left(y_i - \overline{y}
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It will actually help us to rewrite this estimator as

$${\hat eta}_1 = eta_1 + rac{{\sum _i \left({{x_i} - \overline x}
ight){u_i}}}{{\sum _i \left({{x_i} - \overline x}
ight)^2 }}$$

Proof: Assuming $y_i = eta_0 + eta_1 x_i + u_i$

$$egin{aligned} \hat{eta}_1 &= rac{\sum_i \left(y_i - ar{y}
ight) \left(x_i - ar{x}
ight)^2}{\sum_i \left(x_i - ar{x}
ight)^2} \ &= rac{\sum_i \left(\left[eta_0 + eta_1 x_i + u_i
ight] - \left[eta_0 + eta_1 ar{x} + ar{u}
ight]
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ight)}{\sum_i \left(x_i - ar{x}
ight)^2} \end{aligned}$$

$$\begin{split} \hat{\beta}_{1} &= \dots = \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) \left(u_{i} - \overline{u}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \sum_{i} \left(x_{i} - \overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - \sum_{i} \overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - n\overline{x}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i} - \overline{u} \left(\sum_{i} x_{i} - \sum_{i} x_{i}\right)}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \\ &= \beta_{1} + \frac{\sum_{i} \left(x_{i} - \overline{x}\right) u_{i}}{\sum_{i} \left(x_{i} - \overline{x}\right)^{2}} \end{split}$$

Consequences: Bias

We now want to see if heteroskedasticity biases the OLS estimator for β_1 .

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Phew. **OLS is still unbiased** for the β_k .

Consequences: Efficiency

OLS's efficiency and inference do not survive heteroskedasticity.

• In the presence of heteroskedasticity, OLS is **no longer the most efficient** (best) linear unbiased estimator.

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- In the presence of heteroskedasticity, OLS is **no longer the most efficient** (best) linear unbiased estimator.
- It would be more informative (efficient) to weight observations inversely to their u_i 's variance.
 - Downweight high-variance u_i 's (too noisy to learn much).
 - Upweight observations with low-variance u_i 's (more 'trustworthy').
 - Now you have the idea of weighted least squares (WLS)

Consequences: Inference

OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)

Heteroskedasticity

Consequences: Inference

OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)
- It's hard to learn much without sound inference.

Heteroskedasticity

Solutions

- 1. **Tests** to determine whether heteroskedasticity is present.
- 2. **Remedies** for (1) efficiency and (2) inference

While we *might* have solutions for heteroskedasticity, the efficiency of our estimators depends upon whether or not heteroskedasticity is present.

- 1. The Goldfeld-Quandt test
- 2. The Breusch-Pagan test
- 3. The White test

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- 2. The Breusch-Pagan test
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Each of these tests centers on the fact that we can **use the OLS residual** e_i to estimate the population disturbance u_i .

The Goldfeld-Quandt test

Focuses on a specific type of heteroskedasticity: whether the variance of u_i differs **between two groups**.[†]

Remember how we used our residuals to estimate the σ^2 ?

$$s^2 = rac{ ext{SSE}}{n-1} = rac{\sum_i e_i^2}{n-1}$$

We will use this same idea to determine whether there is evidence that our two groups differ in the variances of their disturbances, effectively comparing s_1^2 and s_2^2 from our two groups.

[+]: The G-Q test was one of the early tests of heteroskedasticity (1965).

The Goldfeld-Quandt test

Operationally,

- 1. Order your the observations by \boldsymbol{x}
- 2. Split the data into two groups of size \texttt{n}^{\star}
 - \circ G_1: The first third
 - \circ G_2: The last third
- 3. Run separate regressions of y on x for G_1 and G_2
- 4. Record $\ensuremath{\mathsf{SSE}}_1$ and $\ensuremath{\mathsf{SSE}}_2$
- 5. Calculate the G-Q test statistic

The Goldfeld-Quandt test

The G-Q test statistic

$$F_{(n^{\star}-k,\,n^{\star}-k)} = rac{\mathrm{SSE}_2/(n^{\star}-k)}{\mathrm{SSE}_1/(n^{\star}-k)} = rac{\mathrm{SSE}_2}{\mathrm{SSE}_1}$$

follows an F distribution (under the null hypothesis) with $n^{\star} - k$ and $n^{\star} - k$ degrees of freedom.[†]

The Goldfeld-Quandt test

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Notes

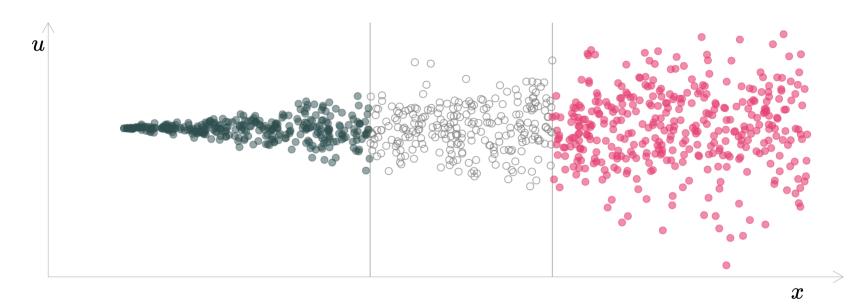
- The G-Q test requires the disturbances follow normal distributions.
- The G-Q assumes a very specific type/form of heteroskedasticity.
- Performs very well if we know the form of potentially heteroskedasticity.

[+]: Goldfeld and Quandt suggested n^* of (3/8)n. k gives number of estimated parameters (*i.e.*, $\hat{\beta}_i$'s).

The Goldfeld-Quandt test



The Goldfeld-Quandt test



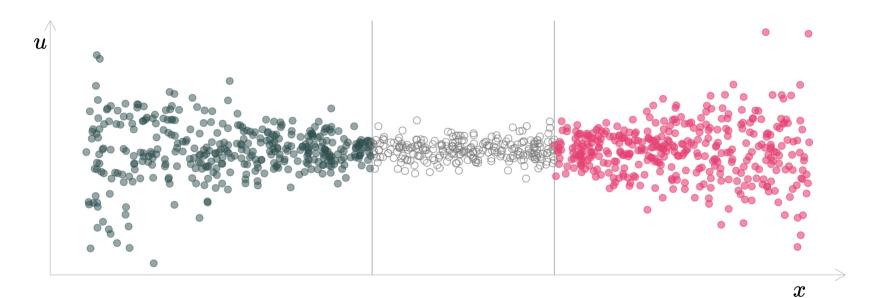
$$F_{375,\,375} = rac{{
m SSE}_2 = 18,\,203.4}{{
m SSE}_1 = 1,\,039.5} pprox 17.5 \implies p ext{-value} < 0.001$$

: We reject H_0 : $\sigma_1^2 = \sigma_2^2$ and conclude there is statistically significant evidence of heteroskedasticity.

The Goldfeld-Quandt test

The problem...

The Goldfeld-Quandt test



$$F_{375,\,375} = rac{\mathrm{SSE}_2 = 14,516.8}{\mathrm{SSE}_1 = 14,937.1} pprox 1 \implies p$$
-value $pprox 0.609$

 \therefore We fail to reject H₀: $\sigma_1^2 = \sigma_2^2$ while heteroskedasticity is present.

The Breusch-Pagan test

Breusch and Pagan (1981) attempted to solve this issue of being too specific with the functional form of the heteroskedasticity.

- Allows the data to show if/how the variance of u_i correlates with X.
- If σ_i^2 correlates with X, then we have heteroskedasticity.
- Regresses e_i^2 on $X = [1, x_1, x_2, \ldots, x_k]$ and tests for joint significance.

The Breusch-Pagan test

How to implement:

1. Regress y on an intercept,
$$x_1$$
, x_2 , ..., x_k .

- 2. Record residuals e.
- 3. Regress e^2 on an intercept, x_1 , x_2 , ..., x_k .

$$e_i^2=lpha_0+lpha_1x_{1i}+lpha_2x_{2i}+\dots+lpha_kx_{ki}+v_i$$

4. Record R^2 .

5. Test hypothesis H $_0$: $lpha_1=lpha_2=\dots=lpha_k=0$

The Breusch-Pagan test

The B-P test statistic[†] is

$${
m LM}=n imes R_e^2$$

where R_e^2 is the R^2 from the regression

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Under the null, LM is asymptotically distributed as χ^2_k .

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This test statistic tests $\mathsf{H}_0: \alpha_1 = \alpha_2 = \cdots = \alpha_k = 0$.

Rejecting the null hypothesis implies evidence of heteroskedasticity.

[+]: This specific form of the test statistic actually comes form Koenker (1981).

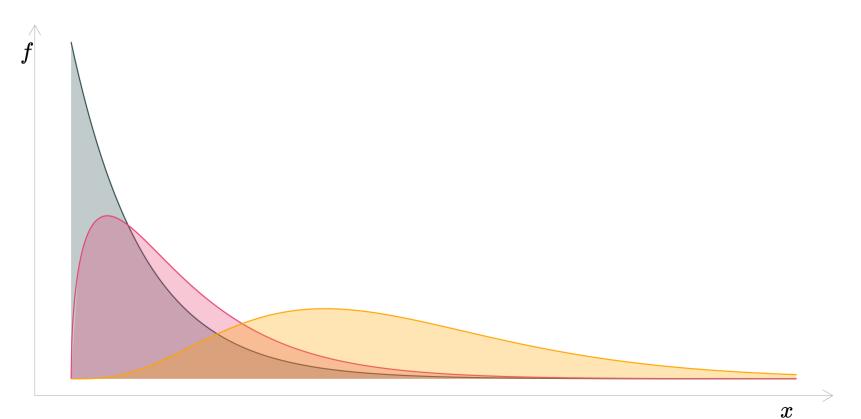
The χ^2 distribution

We just mentioned that under the null, the B-P test statistic is distributed as a χ^2 random variable with k degrees of freedom.

The χ^2 distribution is just another example of a common (named) distribution (like the Normal distribution, the *t* distribution, and the *F*).

The χ^2 distribution

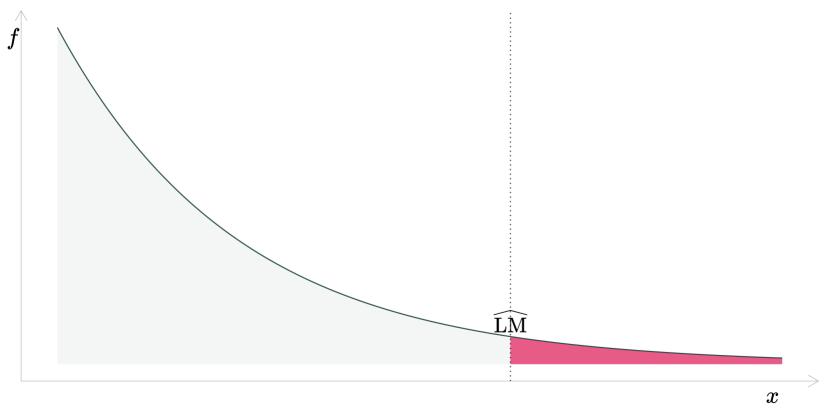
Three examples of χ^2_k : k=1, k=2, and k=9



34 / 66

The χ^2 distribution

Probability of observing a more extreme test statistic $\widehat{\mathbf{LM}}$ under H_0



The Breusch-Pagan test

Problem: We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances σ_i^2 .

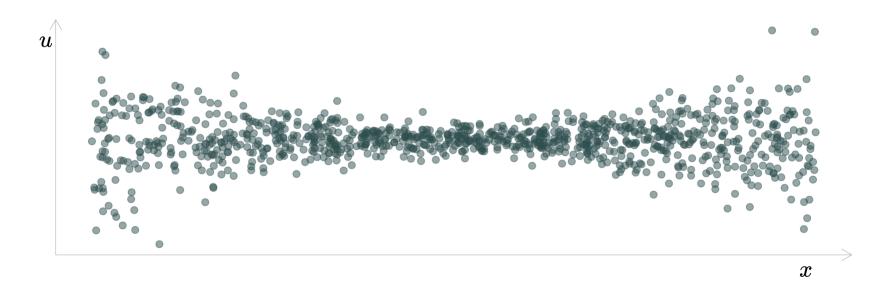
The Breusch-Pagan test

Problem: We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances σ_i^2 .

Result: B-P *may* still miss fairly simple forms of heteroskedasticity.

The Breusch-Pagan test

Breusch-Pagan tests are still **sensitive to functional form**.



$$egin{aligned} e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} & \widehat{ ext{LM}} &= 1.26 & p ext{-value} &pprox 0.261 \ e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} + \hat{lpha}_2 x_{1i}^2 & \widehat{ ext{LM}} &= 185.8 & p ext{-value} &< 0.001 \end{aligned}$$

The White test

So far we've been testing for specific relationships between our explanatory variables and the variances of the disturbances, *e.g.*,

- $\mathsf{H}_0: \sigma_1^2 = \sigma_2^2$ for two groups based upon x_j (**G-Q**)
- $\mathsf{H}_0: \alpha_1 = \cdots = \alpha_k = 0$ from $e_i^2 = \alpha_0 + \alpha_1 x_{1i} + \cdots + \alpha_k x_{ki} + v_i$ (**B-P**)

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However, we actually want to know if

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Q: Can't we just test this hypothesis?

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However, we actually want to know if

$$\sigma_1^2=\sigma_2^2=\dots=\sigma_n^2$$

Q: Can't we just test this hypothesis? **A:** Sort of.

The White test

Toward this goal, Hal White took advantage of the fact that we can **replace the homoskedasticity requirement with a weaker assumption**:

- Old: $\operatorname{Var}(u_i|X) = \sigma^2$
- New: u^2 is uncorrelated with the explanatory variables (*i.e.*, x_j for all j), their squares (*i.e.*, x_j^2), and the first-degree interactions (*i.e.*, x_jx_h).

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This new assumption is easier to explicitly test (*hint*: regression).

The White test

An outline of White's test for heteroskedasticity:

1. Regress y on x_1 , x_2 , ..., x_k . Save residuals e.

2. Regress squared residuals on all explanatory variables, their squares, and interactions.

$$e^2 = lpha_0 + \sum_{h=1}^k lpha_h x_h + \sum_{j=1}^k lpha_{k+j} x_j^2 + \sum_{\ell=1}^{k-1} \sum_{m=\ell+1}^k lpha_{\ell,m} x_\ell x_m + v_i$$

3. Record R_e^2 .

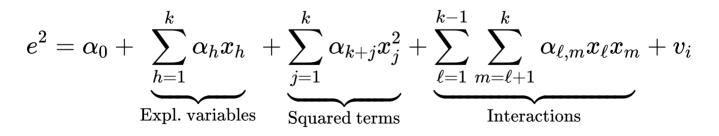
4. Calculate test statistic to test $extsf{H}_0: \; lpha_p = 0$ for all p
eq 0 .

The White test

Just as with the Breusch-Pagan test, White's test statistic is

$$\mathrm{LM} = n imes R_e^2 \qquad \mathrm{Under} \ \mathrm{H}_0, \ \mathrm{LM} \stackrel{\mathrm{d}}{\sim} \chi_k^2$$

but now the R_e^2 comes from the regression of e^2 on the explanatory variables, their squares, and their interactions.



Note: The k (for our χ_k^2) equals the number of estimated parameters in the regression above (the α_j), excluding the intercept (α_0).

The White test

Practical note: If a variable is equal to its square (*e.g.*, binary variables), then you don't (can't) include it. The same rule applies for interactions.

The White test

Example: Consider the model $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u$

Step 1: Estimate the model; obtain residuals (e).

Step 2: Regress e^2 on explanatory variables, squares, and interactions.

$$e^2 = lpha_0 + lpha_1 x_1 + lpha_2 x_2 + lpha_3 x_3 + lpha_4 x_1^2 + lpha_5 x_2^2 + lpha_6 x_3^2 \ + lpha_7 x_1 x_2 + lpha_8 x_1 x_3 + lpha_9 x_2 x_3 + v$$

Record the R^2 from this equation (call it R_e^2).

Step 3: Test H_0 : $lpha_1 = lpha_2 = \dots = lpha_9 = 0$ using $\mathrm{LM} = n R_e^2 \stackrel{\mathrm{d}}{\sim} \chi_9^2$.

[†]: To simplify notation here, I'm dropping the *i* subscripts.

The White test



We've already done the White test for this simple linear regression.

$$e_i^2 = \hatlpha_0 + \hatlpha_1 x_{1i} + \hatlpha_2 x_{1i}^2 \qquad \widehat{ ext{LM}} = 185.8 \qquad p ext{-value} < 0.001$$

Testing for Heteroskedasticity Examples

Examples

Goal: Estimate the relationship between standardized test scores (outcome variable) and (1) student-teacher ratio and (2) income, *i.e.*,

$$(\text{Test score})_i = \beta_0 + \beta_1 \text{Ratio}_i + \beta_2 \text{Income}_i + u_i$$
 (1)

Potential issue: Heteroskedasticity... and we do not observe u_i .

Solution:

- 1. Estimate the relationship in (1) using OLS
- 2. Use the residuals (e_i) to test for heteroskedasticity
 - Goldfeld-Quandt
 - Breusch-Pagan
 - White

Examples

We will use testing data from the dataset Caschool in the Ecdat R package.

Examples

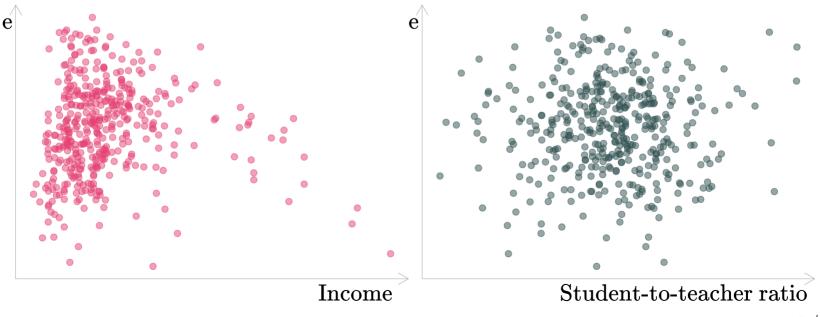
Let's begin by estimating our model

```
(	ext{Test score})_i = eta_0 + eta_1	ext{Ratio}_i + eta_2	ext{Income}_i + u_i
```

#>	#	A tibble: 3	x 5			
#>		term	estimate	std.error	statistic	p.value
#>		<chr></chr>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>
#>	1	(Intercept)	639.	7.45	85.7	5.70e-267
#>	2	ratio	-0.649	0.354	-1.83	6.79e- 2
#>	3	income	1.84	0.0928	19.8	4.38e- 62

Examples

Now, let's see what the residuals suggest about heteroskedasticity



Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

Arrange the data by income
test_df ← arrange(test_df, income)

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Arrange the data by income
test_df ← arrange(test_df, income)
Re-estimate the model for the last and first 158 observations
est_model1 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model2 ← lm(test_score ~ ratio + income, data = head(test_df, 158))

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Grab the residuals from each regression
e_model1 ← residuals(est_model1)
e_model2 ← residuals(est_model2)

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est_model1 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model2 ← lm(test_score ~ ratio + income, data = head(test_df, 158))
# Grab the residuals from each regression
e_model1 ← residuals(est_model1)
e_model2 ← residuals(est_model2)
# Calculate SSE for each regression
(sse model1 ← sum(e model1^2))
```

#> [1] 19305.01

```
(sse_model2 ← sum(e_model2<sup>2</sup>))
```

Example: Goldfeld-Quandt

Remember the Goldfeld-Quandt test statistic?

$$F_{n^{\star}-k,\,n^{\star}-k}=rac{\mathrm{SSE}_2}{\mathrm{SSE}_1}$$

Example: Goldfeld-Quandt

Remember the Goldfeld-Quandt test statistic?

 $F_{n^{\star}-k,\ n^{\star}-k} = rac{\mathrm{SSE}_2}{\mathrm{SSE}_1} pprox rac{29,537.83}{19,305.01}$

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G-Q test statistic
(f_gq ← sse_model2/sse_model1)

#> [1] 1.530061

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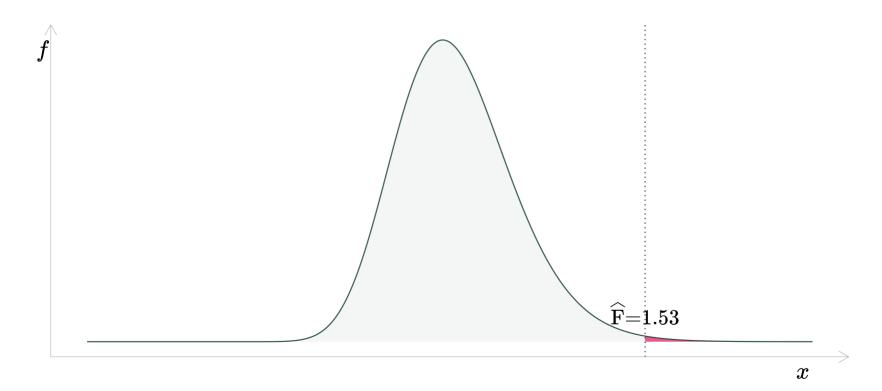
#> [1] 1.530061

p-value
pf(q = f_gq, df1 = 158-3, df2 = 158-3, lower.tail = F)

#> [1] 0.004226666

Example: Goldfeld-Quandt

The Goldfeld-Quandt test statistic and its null distribution



Example: Goldfeld-Quandt

Putting it all together:

 $extsf{H}_{0}\!\!:\sigma_{1}^{2}=\sigma_{2}^{2}$ vs. $extsf{H}_{ extsf{A}}\!\!:\sigma_{1}^{2}
eq\sigma_{2}^{2}$

Goldfeld-Quandt test statistic: Fpprox 1.53

 $p ext{-value} pprox 0.00423$

 \therefore Reject H₀ (*p*-value is less than 0.05).

Conclusion: There is statistically significant evidence that $\sigma_1^2 \neq \sigma_2^2$. Therefore, we find statistically significant evidence of heteroskedasticity (at the 5-percent level).

Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

```
# Arrange the data by ratio
test_df ← arrange(test_df, ratio)
# Re-estimate the model for the last and first 158 observations
est_model3 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model4 ← lm(test_score ~ ratio + income, data = head(test_df, 158))
# Grab the residuals from each regression
e_model3 ← residuals(est_model3)
e_model4 ← residuals(est_model4)
# Calculate SSE for each regression
(sse model3 ← sum(e model3^2))
```

#> [1] 26243.52

```
(sse_model4 ← sum(e_model4<sup>2</sup>))
```

Example: Goldfeld-Quandt

 $F_{n^{\star}-k,\,n^{\star}-k} = rac{\mathrm{SSE}_4}{\mathrm{SSE}_3} pprox rac{29,101.52}{26,243.52} pprox 1.11$

which has a *p*-value of approximately 0.2603.

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Lesson: Understand the limitations of estimators, tests, *etc.*

Example: Breusch-Pagan

Let's test the same model with the Breusch Pagan.

Recall: We saved our residuals as e in our dataset, i.e.,

test_df\$e ← residuals(est_model)

Example: Breusch-Pagan

In B-P, we first regress e_i^2 on the explanatory variables,

Example: Breusch-Pagan

and use the resulting R^2 to calculate a test statistic.

```
# Regress squared residuals on explanatory variables
bp_model ← lm(I(e^2) ~ ratio + income, data = test_df)
# Grab the R-squared
(bp_r2 ← summary(bp_model)$r.squared)
```

#> [1] 3.23205e-05

Example: Breusch-Pagan

The Breusch-Pagan test statistic is

 ${
m LM}=n imes R_e^2$

Example: Breusch-Pagan

- The Breusch-Pagan test statistic is
- ${
 m LM}=n imes R_e^2pprox 420 imes 0.0000323$

Example: Breusch-Pagan

```
The Breusch-Pagan test statistic is
```

 ${
m LM}=n imes R_e^2pprox 420 imes 0.0000323pprox 0.0136$

which we test against a χ^2_k distribution (here: k=2).[†]

Example: Breusch-Pagan

```
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```

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m LM}=n imes R_e^2pprox 420 imes 0.0000323pprox 0.0136$

which we test against a χ^2_k distribution (here: k=2).⁺

#> [1] 0.9932357

[+]: k is the number of explanatory variables (excluding the intercept).

Example: Breusch-Pagan

 $\mathsf{H}_{0}\!\!:lpha_{1}=lpha_{2}=0$ vs. $\mathsf{H}_{\mathsf{A}}\!\!:lpha_{1}
eq 0$ and/or $lpha_{2}
eq 0$

for the model $u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$

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Breusch-Pagan test statistic: $\widehat{LM}\approx 0.014$

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 $p ext{-value} pprox 0.993$

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 \therefore Fail to reject H₀ (the *p*-value is greater than 0.05)

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Conclusion: We do not find statistically significant evidence of heteroskedasticity at the 5-percent level.

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Breusch-Pagan test statistic: $\widehat{LM}\approx 0.014$

 $p ext{-value} pprox 0.993$

 \therefore Fail to reject H₀ (the *p*-value is greater than 0.05)

Conclusion: We do not find statistically significant evidence of heteroskedasticity at the 5-percent level. (We find no evidence of a *linear* relationship between u_i^2 and the explanatory variables.)

Example: Breusch-Pagan

The Breusch-Pagan test statistic and its null distribution



Heteroskedasticity

Example: White

The White test adds squared terms and interactions to the B-P test.

$$egin{aligned} &u_i^2 =& lpha_0 + lpha_1 ext{Ratio}_i + lpha_2 ext{Income}_i \ &+ lpha_3 ext{Ratio}_i^2 + lpha_4 ext{Income}_i^2 + lpha_5 ext{Ratio}_i imes ext{Income}_i \ &+ w_i \end{aligned}$$

which moves the null hypothesis from H₀: $\alpha_1 = \alpha_2 = 0$ to H₀: $\alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = \alpha_5 = 0$

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which moves the null hypothesis from H₀: $\alpha_1 = \alpha_2 = 0$ to H₀: $\alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = \alpha_5 = 0$

So we just need to update our R code, and we're set.

Heteroskedasticity

Example: White

Aside: R has funky notation for squared terms and interactions in lm():

- **Squared terms** use I(), *e.g.*, $lm(y ~ I(x^2))$
- Interactions use : between the variables, *e.g.*, lm(y ~ x1:x2)

Example: Regress y on quadratic of x1 and x2:

```
# Pretend quadratic regression w/ interactions
lm(y ~ x1 + x2 + I(x1<sup>2</sup>) + I(x2<sup>2</sup>) + x1:x2, data = pretend_df)
```

Example: White

Step 1: Regress e_i^2 on 1st degree, 2nd degree, and interactions

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

Example: White

Step 2: Collect R_e^2 from the regression.

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

#> [1] 0.07332222

Example: White

Step 3: Calculate White test statistic $\mathrm{LM} = n imes R_e^2 pprox 420 imes 0.073$

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
white_r2 ← summary(white_model)$r.squared
# Calculate the White test statistic
(white stat ← 420 * white r2)
```

#> [1] 30.79533

Example: White

Step 4: Calculate the associated *p*-value (where $\operatorname{LM} \stackrel{d}{\sim} \chi^2_k$); here, k=5

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
white_r2 ← summary(white_model)$r.squared
# Calculate the White test statistic
white_stat ← 420 * white_r2
# Calculate the p-value
pchisq(q = white_stat, df = 5, lower.tail = F)
```

#> [1] 1.028039e-05

Example: White

Putting everything together...

Example: White

Putting everything together...

 $oxdota_0: lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$

Example: White

Putting everything together...

 $extsf{H}_0: lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ vs. $extsf{H}_{ extsf{A}}: lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$

Example: White

Putting everything together...

 $extsf{H}_0: lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ vs. $extsf{H}_{ extsf{A}}: lpha_i
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$$egin{aligned} u_i^2 =& lpha_0 + lpha_1 ext{Ratio}_i + lpha_2 ext{Income}_i \ &+ lpha_3 ext{Ratio}_i^2 + lpha_4 ext{Income}_i^2 \ &+ lpha_5 ext{Ratio}_i imes ext{Income}_i + w_i \end{aligned}$$

Example: White

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Our White test statistic: ${
m LM}=n imes R_e^2pprox 420 imes 0.073pprox 30.8$

Example: White

Putting everything together...

 $extsf{H}_0: lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ vs. $extsf{H}_{ extsf{A}}: lpha_i
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Our White test statistic: ${
m LM}=n imes R_e^2pprox 420 imes 0.073pprox 30.8$

Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a *p*-value less than 0.001.

Example: White

Putting everything together...

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Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a *p*-value less than 0.001.

: We reject H₀

Example: White

Putting everything together...

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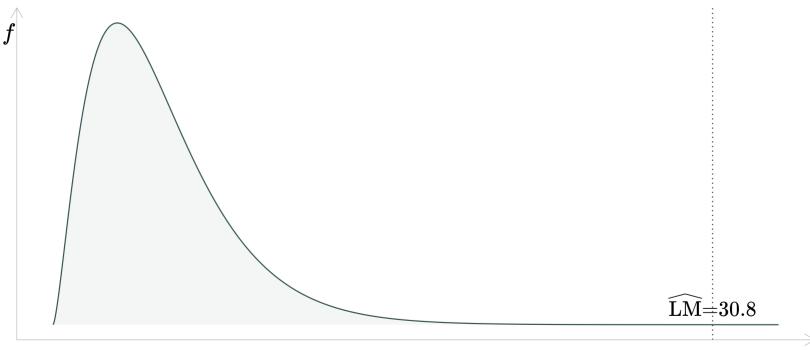
Our White test statistic: ${
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Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a *p*-value less than 0.001.

∴ We **reject H**⁰ and conclude there is **statistically significant evidence of heteroskedasticity** (at the 5-percent level).

Example: White

The White test statistic and its null distribution



- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting *y* against *x*, tell us anything about heteroskedasticity?
- **Q:** Does plotting *e* against *x*, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- **Q:** Which test do you recommend to test for heteroskedasticity? Why?

Review questions

• **Q:** What is the definition of heteroskedasticity?

Review questions

- **Q:** What is the definition of heteroskedasticity?
- A:

Math: $\operatorname{Var}(u_i|X) \neq \operatorname{Var}(u_j|X)$ for some $i \neq j$.

Words: There is a systematic relationship between the variance of u_i and our explanatory variables.

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- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- A: It biases our standard errors—wrecking our statistical tests and confidence intervals. Also: OLS is no longer the most efficient (best) linear unbiased estimator.

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- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- A: It's not exactly what we want, but since y is a function of x and u, it can still be informative. If y becomes more/less disperse as x changes, we likely have heteroskedasticity.

- **Q:** What is the definition of heteroskedasticity?
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- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?
- **A:** Yes. The spread of *e* depicts its variance—and tells us something about the variance of *u*. Trends in this variance, along *x*, suggest heteroskedasticity.

- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?
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- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting *e* against *x*, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- A: We use the e_i 's to predict/learn about the u_i 's. This trick is key for almost everything we do with heteroskedasticity testing/correction.

- **Q:** What is the definition of heteroskedasticity?
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- **Q:** Which test do you recommend to test for heteroskedasticity? Why?

- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- **Q:** Which test do you recommend to test for heteroskedasticity? Why?
- A: I like White. Fewer assumptions. Fewer issues.